

Q.1. A study measured Total Mercury levels (Y, in mg/g) in a sample of n=135 Kuwaiti men. The independent variables were: X<sub>1</sub>=1 if fisherman, 0 if not; X<sub>2</sub> = Weight (kg); and X<sub>3</sub> = # Fish Meals/Week. The matrix results are given below.

X'X					X'Y
135	100	9876	881		509.666
100	100	7280	845		418.083
9876	7280	728452	64639		38360.354
881	845	64639	9529		3959.497
INV(X'X)					beta-hat
0.967339	-0.060848	-0.012685	0.002005		-11.064
-0.060848	0.063336	0.000480	-0.003248		1.027
-0.012685	0.000480	0.000171	-0.000033		0.183
0.002005	-0.003248	-0.000033	0.000432		0.106
Y'Y					
3081.235					

$$n = 135$$

$$p = 3 \quad p' = 4$$

$$\sum_i Y_i = 509.66$$

$$\Rightarrow \frac{(\sum Y_i)^2}{n} = \frac{509.66^2}{135} = 1924.10$$

p.1.a Complete the following Analysis of Variance table.

$$SS_{TOTALC} = 3081.235 - 1924.10 = 1157.14$$

Source	df	SS	MS	F_obs	F(0.05)
Regression	3	305.975	101.992	15.698	≈ 2.674
Residual	131	851.16	6.497	#N/A	#N/A
Total (Corr)	134	1157.14	#N/A	#N/A	#N/A

$$Y'PY = \hat{\beta}'X'Y = -11.064(509.666) + 1.027(418.083) + 0.183(38360.354) + 0.106(3959.497)$$

$$= -5638.945 + 429.371 + 7019.945 + 419.704 = 2230.075$$

$$Y'Y - Y'PY = 3081.235 - 2230.075 = 851.16$$

$$Y'(p - \frac{1}{n}J)Y = 2230.075 - 1924.10 = 305.975$$

p.1.b. Obtain a 95% Confidence Interval for the effect of being a fisherman on expected total Mercury, controlling for Weight and Fish Meals/Week.

$$\text{CI} = 1.027 \pm 1.978 \sqrt{6.497(.063336)} = 1.027 \pm 1.978(.6415)$$

$$= 1.027 \pm 1.2688$$

$$= (-.242, 2.296)$$

p.1.c. What proportion of the variance in Total Mercury is "explained" by this set of predictors?

$$R^2 = \frac{305.975}{1157.14} = 0.2644$$

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Q.2. For the multiple regression model, with an intercept term, complete the following parts, **showing all of your work.**

NOTE:  $[(X'X)^{-1}]' = (X'X)^{-1}$

p.2.a.  $Y = \hat{Y} + e$      $\hat{Y} = X\hat{\beta}$     prove that  $Y'Y = \hat{Y}'\hat{Y} + e'e$

$$\hat{Y} = X\hat{\beta} = X(X'X)^{-1}X'Y = PY$$

(15)  $P' = (X(X'X)^{-1}X')' = X(X'X)^{-1}X' = P$      $PP = X(X'X)^{-1}X'X(X'X)^{-1}X' = X(X'X)^{-1}X' = P$

$$\hat{Y}'\hat{Y} = (PY)'(PY) = Y'P'PY = Y'PPY = Y'PY$$

$$e = Y - \hat{Y} = (I - P)Y \quad (I - P)' = I' - P' = I - P$$

$$(I - P)(I - P) = I - P - P + PP = I - P - P + P = I - P$$

$$\Rightarrow e'e = [(I - P)Y]'(I - P)Y = Y'(I - P)'(I - P)Y = Y'(I - P)(I - P)Y = Y'(I - P)Y$$

$$Y'Y = Y'IY = Y'PY + Y'(I - P)Y = \hat{Y}'\hat{Y} + e'e \quad \checkmark$$

p.2.b. Derive the sampling distributions of  $\hat{Y}$  and  $e$

(15)  $\hat{Y} = PY \sim \text{Normal w/}$

$$E\{\hat{Y}\} = PE\{Y\} = PX\beta = X(X'X)^{-1}X'X\beta = X\beta$$

$$V\{\hat{Y}\} = P\sigma^2IP' = \sigma^2PP' = \sigma^2PP = \sigma^2P$$

$$\hat{Y} \sim N(X\beta, \sigma^2P)$$

$$e = (I - P)Y \sim \text{Normal w/}$$

$$E\{e\} = (I - P)E\{Y\} = (I - P)X\beta = X\beta - PX\beta = X\beta - X\beta = 0$$

$$V\{e\} = V\{(I - P)Y\} = (I - P)\sigma^2I(I - P)' = \sigma^2(I - P)(I - P)'$$

$$= \sigma^2(I - P)(I - P) = \sigma^2(I - P) \quad e \sim N(0, \sigma^2(I - P))$$

p.2.c. Derive the sampling distributions of  $\frac{1}{\sigma^2} \hat{\mathbf{Y}}' \hat{\mathbf{Y}}$  and  $\frac{1}{\sigma^2} \mathbf{e}' \mathbf{e}$

$$(15) \hat{\mathbf{Y}}' \hat{\mathbf{Y}} = \mathbf{Y}' \mathbf{P} \mathbf{Y} \quad \mathbf{P} \mathbf{I} \mathbf{P} \mathbf{I} = \mathbf{P} \mathbf{P} = \mathbf{P} = \mathbf{P} \mathbf{I}$$

$$\Rightarrow \frac{1}{\sigma^2} \hat{\mathbf{Y}}' \hat{\mathbf{Y}} \sim \chi^2 \text{ w/ } df = \text{rank}(\mathbf{P}) = \text{tr}(\mathbf{P}) = \text{tr}(\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}') \\ = \text{tr}(\mathbf{X}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}) = \text{tr}(\mathbf{I}_{p'}) = p'$$

$$\mathbf{Q} = \frac{1}{2\sigma^2} \beta' \mathbf{X}' \mathbf{P} \mathbf{X} \beta = \frac{1}{2\sigma^2} \beta' \mathbf{X}' \mathbf{X} \beta \quad (\mathbf{P} \mathbf{X} = \mathbf{X})$$

$$\Rightarrow \frac{1}{\sigma^2} \hat{\mathbf{Y}}' \hat{\mathbf{Y}} \sim \chi^2(p', \frac{1}{2\sigma^2} \beta' \mathbf{X}' \mathbf{X} \beta)$$

$$\mathbf{e}' \mathbf{e} = \mathbf{Y}' (\mathbf{I} - \mathbf{P}) \mathbf{Y} \quad (\mathbf{I} - \mathbf{P}) \mathbf{I} (\mathbf{I} - \mathbf{P}) \mathbf{I} = (\mathbf{I} - \mathbf{P}) \mathbf{I}$$

$$\Rightarrow \frac{1}{\sigma^2} \mathbf{e}' \mathbf{e} \sim \chi^2 \text{ w/ } df = \text{rank}(\mathbf{I} - \mathbf{P}) = \text{trace}(\mathbf{I} - \mathbf{P}) = n - p'$$

$$\mathbf{Q} = \frac{1}{2\sigma^2} \beta' \mathbf{X}' (\mathbf{I} - \mathbf{P}) \mathbf{X} \beta = 0 \quad (\mathbf{I} \mathbf{X} \beta = \mathbf{P} \mathbf{X} \beta = \mathbf{X} \beta) \Rightarrow \frac{1}{\sigma^2} \mathbf{e}' \mathbf{e} \sim \chi^2_{n-p'}$$

p.2.d. Show that  $\frac{1}{\sigma^2} \hat{\mathbf{Y}}' \hat{\mathbf{Y}}$  and  $\frac{1}{\sigma^2} \mathbf{e}' \mathbf{e}$  are independent.

$$(8) \mathbf{A} = \mathbf{P} \quad \mathbf{V} = \mathbf{I} \quad \mathbf{B} = (\mathbf{I} - \mathbf{P})$$

$$\mathbf{A} \mathbf{V} \mathbf{B} = \mathbf{P} \mathbf{I} (\mathbf{I} - \mathbf{P}) = \mathbf{P} (\mathbf{I} - \mathbf{P}) = \mathbf{P} - \mathbf{P} \mathbf{P} = \mathbf{P} - \mathbf{P} = \mathbf{0}$$

Q.3. A forensic study related Hand ( $X_1$ ) and Foot ( $X_2$ ) lengths to Height ( $Y$ ) for a sample of  $n = 75$  adult females (each variable in 100s of mms). Consider the following three models.

$$M_1: E\{Y\} = \beta_0 + \beta_1 X_1 \quad M_2: E\{Y\} = \beta_0 + \beta_2 X_2 \quad M_3: E\{Y\} = \beta_0 + \beta_1 X_1 + \beta_2 X_2$$

Model 1 (Hand)			Model 2 (Foot)		
X'X		X'Y	X'X		X'Y
75	142.185	1199.70	75	176.062	1199.70
142.185	270.1992	2276.80	176.062	414.390402	2819.37
INV(X'X)		beta-hat	INV(X'X)		beta-hat
5.586829	-2.93992	8.9042	5.087522	-2.16153487	9.3261
-2.93992	1.550753	3.7408	-2.161535	0.920784244	2.8413
Y'Y	Model 3 (Hand, Foot)				
19208.28	X'X		X'Y		
	75	142.185	176.062	1199.70	
	142.185	270.1992	334.2884	2276.80	
	176.062	334.2884	414.3904	2819.37	
	INV(X'X)		beta-hat		
	6.640061	-1.95728	-1.24223	7.4414	
	-1.95728	2.467531	-1.15897	2.3760	
	-1.24223	-1.15897	1.465136	1.7253	

p.3.a. Compute  $Y' \left( \frac{1}{n} \mathbf{J} \right) Y$  and the Total (Corrected) Sum of Squares.

$$(6) \quad Y' \frac{1}{n} \mathbf{J} Y = \frac{1}{n} \left( \sum_i Y_i \right)^2 = \frac{1}{75} (1199.70)^2 = 19190.40$$

$$(6) \quad SS_{TOTALc} = Y' (I - \frac{1}{n} \mathbf{J}) Y = 19208.28 - 19190.40 = 17.88$$

p.3.b. Compute the Residual (Error) Sum of Squares for each model. 3.

$$(5) \quad Y' P_1 Y = 8.9042 (1199.70) + 3.7408 (2276.80) = 19199.42$$

$$(5) \quad Y' P_2 Y = 9.3261 (1199.70) + 2.8413 (2819.37) = 19199.20$$

$$(5) \quad Y' P_3 Y = 7.4414 (1199.70) + 2.3760 (2276.80) + 1.7253 (2819.37) = 19201.38$$

$$SSE_1 = 19208.28 - 19199.42 = 8.86$$

$$SSE_2 = 19208.28 - 19199.20 = 9.08$$

$$(6) \quad SSE_3 = 19208.28 - 19201.38 = 6.90$$

p.3.c. Compute  $R(\beta_1 | \beta_0), R(\beta_2 | \beta_0), R(\beta_1 | \beta_0, \beta_2), R(\beta_2 | \beta_0, \beta_1)$

$$R(\beta_1 | \beta_0) = 19199.42 - 19190.40 = 9.02 \quad (5)$$

$$R(\beta_2 | \beta_0) = 19199.20 - 19190.40 = 8.80 \quad (5)$$

$$R(\beta_1 | \beta_0, \beta_2) = 19201.38 - 19199.20 = 2.18 \quad (5)$$

$$R(\beta_2 | \beta_0, \beta_1) = 19201.38 - 19199.42 = 1.96 \quad (5)$$

p.3.d. Use the general linear test for Model 3 to test  $H_0: \beta_1 - \beta_2 = 0$  vs  $H_A: \beta_1 - \beta_2 \neq 0$

$$K' = [0 \quad 1 \quad -1 \quad 0] \quad K'\hat{\beta} = 2.3760 - 1.7253 = 0.6507$$

$$K'(X'X)^{-1}K = 2.467531 + 1.465136 + 2(-1.15897) = \cancel{6.2506} \quad 6.2506$$

$$\Rightarrow Q = \frac{(0.6507)^2}{\cancel{6.2506}} = \cancel{0.0677} \quad 0.0677 \quad k=1$$

$$S^2 = MSE_3 = \frac{6.90}{75-3} = 0.0958 \quad (5)$$

$$T.S. F_{obs} = \frac{Q/k}{S^2} = \frac{\cancel{0.0677}}{0.0958} = \cancel{0.7071} \quad 0.7071 \quad (3)$$

RR:  $F_{0.05, 1, 72} \approx \cancel{3.975} \quad 3.975$  Fail to reject  $H_0$   
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Q.4. A simple linear regression model was fit, relating Weight (Y, in pounds) to Height (X, in inches) among a sample of n=9 Women's NBA basketball players. There were 3 distinct height levels, with 3 players per height. Use the tables below to conduct the F-test for Lack of Fit.

$H_0: E\{Y_{ij}\} = \beta_0 + \beta_1 X_j$  vs  $H_A: E\{Y_{ij}\} = \mu_j \neq \beta_0 + \beta_1 X_j$

$C=3$   $P'=2$   $n-C=6$   
 $C-P'=1$

Height	Weight	Ybar(Grp)	Yhat(Grp)	PE	LF		
68	130	150	145.83	$(-20)^2 = 400$	$4.17^2 = 17.39$		Coefficients Intercept -321.667 Height 6.875
68	165	150	145.83	225	17.39		
68	155	150	145.83	25	17.39		
72	180	165	173.33	225	69.39		
72	165	165	173.33	0	69.39	$\hat{Y}_{68} = 145.83$	
72	150	165	173.33	225	69.39	$\hat{Y}_{72} = 173.33$	
76	200	205	200.83	25	17.39	$\hat{Y}_{76} = 200.83$	
76	185	205	200.83	400	17.39		
76	230	205	200.83	625	17.39		

2150      312.51  
 3 each      9 each      3 each      3      3

Source	df	SS	MS	F	F(.05)
Lack of Fit	1	312.51	312.51	0.87	5.987
Pure Error	6	2150	358.33	#N/A	#N/A
Error	7	2462.51	351.79	#N/A	#N/A

Q.5. A linear regression model is fit, relating mean January temperatures (Y, in °F) to Elevation (X<sub>1</sub>, in 100s of feet) and Latitude (X<sub>2</sub>, in degrees north latitude) for a random sample of n = 63 weather stations in Texas. The (partial) computer results are given below.

n = 63    p = 3    p = 2

ANOVA	2 each		3 each	3	3
	df	SS	MS	F	F(.05)
Regression	2	2472.2	1236.1	807.03	3.150
Residual	60	91.9	1.53	#N/A	#N/A
Total	62	2564.1	#N/A	#N/A	#N/A
<b>Coefficients</b>					
	<b>Standard Error</b>	<b>t Stat</b>	<b>P-value</b>		
Intercept	115.906	2.478	46.776	0.0000	
ELEV.C	-0.117	0.013	-8.877	0.0000	
LAT	-2.183	0.082	-26.580	0.0000	

p.5.a. Complete the ANOVA table.

p.5.b. Dallas/Fort Worth International Airport (DFW) was not in the sampled locations, and is located at an elevation of X<sub>1</sub> = 5.6 and a latitude of X<sub>2</sub> = 32.9. Give the predicted value for DFW.

$$\hat{Y}_{DFW} = 115.906 - 0.117(5.6) + 2.183(32.9) = 43.43 \quad (5)$$

p.5.c. For DFW, we obtain the following values:  $x_0 = \begin{bmatrix} 1 \\ 5.6 \\ 32.9 \end{bmatrix}$      $x_0'(X'X)^{-1}x_0 = 0.0454$ .

Compute the 95% **Prediction Interval** for DFW's mean January temperature.     $t_{.025, 60} = 2.000$

$$43.43 \pm 2.000 \sqrt{1.53(1 + 0.0454)} = 43.43 \pm 2.000(1.2647) \\ = (40.90, 45.96) \quad (3)$$