## Experimental Design Problems

## Part A: 1-Way ANOVA (Completelely Randomized Design)

QA.1. An experiment was conducted as a Completely Randomized Design (1-Way ANOVA) to compare $t=4$ methods of packaging steaks, in terms of the amount of bacteria measured after 9 days of storage. There were $n_{i}=3$ replicates per treatment. The treatment means and sums of squares were:

$$
\bar{Y}_{1 \bullet}=7.48 \quad \bar{Y}_{2 \bullet}=5.50 \quad \bar{Y}_{3 \bullet}=7.26 \quad \bar{Y}_{4 \bullet}=3.36 \quad \sum_{i=1}^{4} n_{i}\left(\bar{Y}_{i \bullet}-\bar{Y}_{\bullet \bullet}\right)^{2}=32.87 \quad \sum_{i=1}^{4}\left(n_{i}-1\right) s_{i}^{2}=0.93
$$

p.1.a. Conduct the F-test for testing $H_{0}: \mu_{1}=\mu_{2}=\mu_{3}=\mu_{4} \quad\left(\alpha_{1}=\alpha_{2}=\alpha_{3}=\alpha_{4}=0\right)$

Test Statistic: $\qquad$ Rejection Region: $\qquad$ Reject $\mathrm{H}_{0}$ ? Yes or No
p.1.b. Compute Tukey's Honest Significant Difference for simultaneously comparing all pairs of packages, with a family-wise error rate of 0.05 . Identify significant differences among all pairs of means.

$$
\begin{array}{cccc}
\text { Trt4 } & \text { Trt2 } & \text { Trt3 } & \text { Trt1 }
\end{array}
$$

p.1.c. Compute Bonferroni's Minimum Significant Difference for simultaneously comparing all pairs of packages, with a family-wise error rate of 0.05

$$
\begin{array}{llll}
\text { Trt4 } & \text { Trt2 } & \text { Trt3 } & \text { Trt1 }
\end{array}
$$

QA.2. A 1-Way ANOVA is conducted to compare the effects of 4 methods of preparing steel. Five replicates of each method are obtained, and the breaking strength is measured. Suppose that the between treatment sum of squares is 1200 , and the within treatment sum of squares is 2400 . Give the test statistic for testing whether the true mean breaking strengths differ among the 4 methods. Give the minimum significant difference for pairs of methods, based on Bonferroni's method with an experimentwise error rate of 0.05 .

QA.3. For a 1-Way ANOVA, based on 3 treatments, and 30 subjects per treatment, give the Treatment and Error Degrees of Freedom:
$\mathrm{Df}_{\mathrm{Trt}}=$ $\qquad$

$$
\mathrm{df}_{\mathrm{Err}}=
$$

$\qquad$

QA.4. A Completely Randomized Design is conducted to compare 5 varieties of fertilizer on plant yield. Each variety is randomly assigned to 7 plots of land (each plot only receives one variety).

DF(Treatments) $\qquad$ DF(Error) DF(Total) $\qquad$

QA.5. When using Bonferroni's method of adjustment for simultaneous Confidence Intervals, as the number of intervals increases, the width of the individual confidence intervals will decrease. $\qquad$

QA.6. An experiment is run to compare $\mathrm{t}=4$ meat packaging conditions. There were $\mathrm{n}_{\mathrm{i}}=3$ replicates per treatment in the Completely Randomized Design. The response was a measure of bacteria count (high values are bad). The treatment means and standard deviations are given below for the model: $\mathrm{Y}_{\mathrm{ij}}=\mu_{\mathrm{i}}+\varepsilon_{\mathrm{ij}}$.
p.6.a. Compute the Treatment and Error Sum of Squares:

| Treatment | Mean | SD | SS(Treatments) | SS(Error) |
| :---: | :---: | :---: | :---: | :---: |
| 1 | 7.48 | 0.44 |  |  |
| 2 | 5.50 | 0.27 |  |  |
| 3 | 7.26 | 0.19 |  |  |
| 4 | 3.36 | 0.40 |  |  |
| Overall | 5.90 | Total |  |  |

p.6.b. Compute the F-Statistic for testing $\mathrm{H}_{0}: \mu_{1}=\mu_{2}=\mu_{3}=\mu_{4}$
p.6.c. Conclude packaging condition true means not all equal if test statistic falls in the range $\qquad$
p.6.d. Based on your test, the P-value will be Larger / Smaller than 0.05

QA.7. An experiment is conducted as a Completely Randomized Design with $t=5$ treatments and $n_{i}=5$ replicates per treatment. The error sum of squares is $\operatorname{SSE}=250$. Compute Bonferroni's minimum significant difference for all pairwise comparisons with experiment-wise error rate of $\alpha_{E}=0.05$.
$\mathrm{B}_{\mathrm{ij}}=$
QA.8. A Completely Randomized Design is conducted with 3 treatments, and 8 replicates per treatment (independent samples). Once the measurements have been ranked from smallest to largest, adjusting for ties, you compute the rank sums to be: $\mathrm{T}_{1}=110, \mathrm{~T}_{2}=100, \mathrm{~T}_{3}=90$. You conduct the Kruskal-Wallis test, $\alpha=0.05$ :
p.8.a. Test Statistic:
p.8.b. Conclude treatment means (medians) are significantly different if Test Stat falls in range: $\qquad$

QA.9. An experiment was conducted as a Completely Randomized Design (1-Way ANOVA) to compare $t=4$ methods of packaging steaks, in terms of the amount of bacteria measured after 9 days of storage. There were $n_{i}=3$ replicates per treatment. The treatment means and sums of squares were:
$\bar{Y}_{1 \bullet}=7.48 \quad \bar{Y}_{2 \bullet}=5.50 \quad \bar{Y}_{3 \bullet}=7.26 \quad \bar{Y}_{4 \bullet}=3.36 \quad \sum_{i=1}^{4} n_{i}\left(\bar{Y}_{i \bullet}-\bar{Y}_{\bullet \bullet}\right)^{2}=32.87 \quad \sum_{i=1}^{4}\left(n_{i}-1\right) s_{i}^{2}=0.93$
p.9.a. Conduct the F-test for testing $H_{0}: \mu_{1}=\mu_{2}=\mu_{3}=\mu_{4} \quad\left(\alpha_{1}=\alpha_{2}=\alpha_{3}=\alpha_{4}=0\right)$

Test Statistic: $\qquad$ Rejection Region: $\qquad$ Reject $\mathrm{H}_{0}$ ? Yes or No
p.9.b. Compute Tukey's Honest Significant Difference for simultaneously comparing all pairs of packages, with a family-wise error rate of 0.05 . Identify significant differences among all pairs of means.
$\begin{array}{llll}\text { Trt4 } & \text { Trt2 } & \text { Trt3 }\end{array}$
p.9.c. Compute Bonferroni’s Minimum Significant Difference for simultaneously comparing all pairs of packages, with a family-wise error rate of 0.05
Trt4
Trt2
Trt3
Trt1

QA.10. . Researchers studied nest humidity levels among 54 species of birds. The nests were classified as (1=Cup, $2=S c r a p e, 3=C o v e r e d)$. The following table gives the sample sizes, means, and standard deviations among the 3 nest types.
p.10.a. Test whether the population mean nest humidity levels differ among the 3 nest types (first obtain the relevant sums of squares and degrees of freedom). $\mathrm{H}_{0}: \mu_{1}=\mu_{2}=\mu_{3}$

| NestType | n | mean | SD |
| ---: | ---: | ---: | ---: |
| 1 | 39 | 20.84 | 4.76 |
| 2 | 9 | 19.98 | 4.19 |
| 3 | 6 | 31.74 | 3.20 |
| Overall | 54 | 21.91 | \#N/A |


| ANOVA |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Source | df | SS |  | MS | F | F(.05) |
| Nest Type |  |  |  |  | Reject H0? |  |
| Error |  |  |  |  |  |  |
| Total |  |  |  |  |  |  |

p.10.b. Use Bonferroni's method to obtain the minimum significant difference between each pair of means.
$\qquad$ Cup vs Covered: $\qquad$ Scrape vs Covered: $\qquad$

QA.11. A study compared infarct volumes of mice exposed to one of 3 treatments in a completely randomized design ( $1=$ vehicle control, $2=$ compound $\mathrm{X}, 3=$ compound Y ). There were a few extreme outliers, so the Kruskal-Wallis test will be applied. The following table gives the sample sizes and rank sums for the 3 treatments. Conduct the Kruskal-Wallis test to determine whether the population medians differ among the 3 treatments.

| Trt | $\mathbf{N}$ | RankSum |
| :---: | :---: | :---: |
| 1 | 13 | 326 |
| 2 | 15 | 375 |
| 3 | 14 | 202 |

QA.12. A published report, based on a balanced 1-Way ANOVA reports means (SDs) for the three treatments as:
Trt 1: 70 (8) Trt 2: 75 (6) $\operatorname{Trt} 3: 80$ (10)
Unfortunately, the authors fail to give the sample sizes.
p.12.a. Complete the following table, given arbitrary levels of the number of replicates per treatment:

| $r$ | SSTrt | SSErr | MSTrt | MSErr | F_obs | $F(.05)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 |  |  |  |  |  |  |
| 6 |  |  |  |  |  |  |
| 10 |  |  |  |  |  |  |

p.12.b. The smallest r , so that these means are significantly different is:
i) $\mathrm{r}<=2$
ii) $2<r<=6$
iii) $6<r<=10$
iv) $r>10$

QA.13. An experiment is conducted as a Completely Randomized Design to compare the durability of 5 green fabric dyes, with respect to washing. A sample of 30 plain white $t$-shirts was obtained, and randomized so that 6 received each dye (with each shirt receiving exactly one dye). A measure of the color brightness of the shirts after 10 wash/dry cycles is obtained (with higher scores representing brighter color). The error sum of squares is reported to be $\operatorname{SSE}=2000$. The mean scores for the 5 dyes are: $\bar{Y}_{1 \bullet}=30 \quad \bar{Y}_{2 \bullet}=25 \quad \bar{Y}_{3 \bullet}=40 \quad \bar{Y}_{4 \bullet}=35 \quad \bar{Y}_{5 \bullet}=20$
p.13.a. Compute Tukey's HSD, and determine which (if any) pairs of means are significantly different with an experiment-wise (overall) error rate of $\alpha_{E}=0.05$.

Tukey's HSD: $\qquad$
p.13.b. Compute the Bonferroni MSD, and determine which (if any) pairs of means are significantly different with an experiment-wise (overall) error rate of $\alpha_{E}=0.05$.

Bonferroni's MSD: $\qquad$

QA.14. A study compared efficiency levels (based on a complex algorithm) among three types of Trade Shows in Spain. The authors classified Trade Shows as being one of 3 sectors (Consumer Goods, Investment Goods, and Services). The Trade Shows were ranked based on their efficiencies (1=Lowest). Based on the sample sizes and the Rank Sums from the following table, conduct the Kruskal-Wallis Test (Note: Total is NOT a "treatment," it is just useful in computations).

| Sector | n | RankSum |
| :---: | :---: | :---: |
| Consumer | 21 | 466 |
| Investment | 16 | 312 |
| Services | 8 | 257 |
| Total | 45 | 1035 |

Test Statistic: $\qquad$ Rejection Region: $\qquad$

QA.15. A study compared antioxidant activity of $t=8$ brands of craft beer in a 1 -Way ANOVA. One response reported was DPPH radical scavenging activity. Each brand was had $n=3$ replicates measured.

| Brand | Mean | SD |
| :--- | ---: | ---: |
| L | 794.9 | 27.5 |
| P | 376 | 32.4 |
| W | 706.2 | 30.9 |
| B9 | 586.5 | 17.7 |
| N | 864.6 | 42.4 |
| R | 670.6 | 19.9 |
| T | 1310.3 | 19 |
| E | 679.2 | 25 |

$$
\sum_{i=1}^{t}\left(\bar{y}_{i \bullet}-\bar{y}_{. .}\right)^{2}=508911.8 \quad \sum_{i=1}^{t} s_{i}^{2}=6253.88
$$

p.15.a. Complete the following Analysis of Variance table used to test $\mathrm{H}_{0}: \mu_{1}=\ldots .=\mu_{8}$

| Source | df | SS | MS | F_obs | F(.05) |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Brand |  |  |  |  |  |
| Error |  |  |  |  |  |
| Total |  |  |  |  |  |

p.15.b. Do we reject the null hypothesis, and conclude the population means differ among the brands? Yes or No
p.15.c. Compute Tukey's minimum significant difference and determine which brands are significantly different.
P
B9
R
E
W
L
T

QA.16. A study classified a sample of French Ski resorts into 3 classifications (large, medium, and small) based on their volume of business. The researchers obtained a measure of each resort's Luenberger Productivity Index (LPI) was obtained. The authors conducted a Kruskal-Wallis test to test whether population median LPI scores differ by resort size group. The numbers and rank sums for each resort size group are given below.

| Size | n | RankSum |
| :--- | ---: | ---: |
| Large | 16 | 428 |
| Medium | 31 | 932 |
| Small | 17 | 720 |

Test Statistic: $\qquad$ Rejection Region $\qquad$ P -value is $>\mathbf{0 . 0 5}$ or $<\mathbf{0 . 0 5}$

QA.17. An experiment was conducted to determine the effect of $g=3$ different food portion/container sizes on food intake in a Completely Randomized Design. There were a total of $N=90$ subjects who were randomized so that 30 received each condition (each subject was observed in one of the 3 conditions). The conditions were: $1=$ medium portion/small container, $2=$ medium portion/large container, $3=$ large proportion/large continer. The response was food intake ( Y , in grams) that the subject consumed while watching a television show. The model and summary statistics are given below.
$y_{i j}=\mu+\alpha_{i}+\varepsilon_{i j}=\mu_{i}+\varepsilon_{i j} \quad n_{1}=30, \bar{y}_{1 \bullet}=30, s_{1}=30 \quad n_{2}=30, \bar{y}_{2 \bullet}=69, s_{2}=44 \quad n_{3}=30, \bar{y}_{3 \bullet}=60, s_{3}=45$
p.17.a. Compute the Between Treatment Sum of Squares (SST) and Within Treatment Sum of Squares (SSE).

SST $=$ $\qquad$ $\mathrm{SSE}=$ $\qquad$
p.17.b. Test $H_{0}: \alpha_{1}=\alpha_{2}=\alpha_{3}=0$

Test Statistic: $\qquad$ Rejection Region $\qquad$ P -value $>$ or $<0.05$
p.17.c. Use Tukey's method to compare all pairs of treatments.

Tukey's $\mathrm{W}=$ $\qquad$ Trt1 Trt3 Trt2

QA.18. Consider the following 3 scenarios for a (Fixed Effects) Completely Randomized Design.
$y_{i j}=\mu_{i}+\varepsilon_{i j} \quad i=1,2,3 ; j=1, \ldots, n \quad \varepsilon_{i j} \sim N\left(0, \sigma^{2}\right)$
a) $\mu_{1}=80, \mu_{2}=100, \mu_{3}=120, \sigma=20, n=5 \quad$ b) $\mu_{1}=90, \mu_{2}=100, \mu_{3}=110, \sigma=10, n=3 \quad$ c) $\mu_{1}=95, \mu_{2}=100, \mu_{3}=105, \sigma=5, n=7$

Rank the from smallest to largest in terms of $\frac{E\{M S T\}}{E\{M S E\}}$
Smallest: $\qquad$ Middle: $\qquad$ Largest: $\qquad$

QA.19. A delivery company is considering buying one of 3 drones for deliveries. They fly each drone 12 times, measuring the distance from the landing point to the target. Due to the skewed distribution of the distances, they use the nonparametric Kruskal-Wallis procedure to test for differences among the drones' true medians. The rank sums are 200, 218, and 248 for the 3 drones. Test $\mathrm{H}_{0}: \mathrm{M}_{1}=\mathrm{M}_{2}=\mathrm{M}_{3}$.

Test Statistic: $\qquad$ Rejection Region: $\qquad$ P-value < or > . 05

QA.20. Unless the number of treatments is 2, Tukey's HSD (W) will always be smaller that Bonferroni's MSD (B) for a given set of data. True / False

QA.21. An experiment was conducted to compare the effects of 4 fragrances on various office workers characteristics. There were 50 subjects per treatment (fragrance). One response measured was the workers' concentration levels. The experiment was conducted as a Completely Randomized Design.

$$
y_{i j}=\mu_{i}+\varepsilon_{i j} \quad i=1, \ldots, 4 ; j=1, \ldots, 50 \quad \varepsilon_{i j} \sim N\left(0, \sigma^{2}\right)
$$

| Trt (i) | $\mathrm{n} \_\mathrm{i}$ | ybar i | $\mathrm{s} \_\mathrm{i}$ |
| :--- | :---: | :---: | :---: |
| Control | 50 | 103.3 | 9.0 |
| Citrus | 50 | 105.2 | 8.8 |
| Grapefruit | 50 | 103.7 | 9.4 |
| Rose | 50 | 104.2 | 9.3 |

p.21.a. Compute the Between treatment sum of squares (SST) and its degrees of freedom ( $\mathrm{df}_{\mathrm{T}}$ )

SST = $\qquad$ $\mathrm{df}_{\mathrm{T}}=$ $\qquad$
p.21.b. Compute the Within treatment sum of squares (SSE) and its degrees of freedom ( $\mathrm{df}_{\mathrm{E}}$ )
$\mathrm{SSE}=$ $\qquad$ $\mathrm{df}_{\mathrm{E}}=$ $\qquad$ p.21.c.Test whether there is evidence of treatment effects. $H_{0}: \mu_{1}=\mu_{2}=\mu_{3}=\mu_{4} \quad H_{A}:$ Not all $\mu_{i}$ are equal Test Statistic: $\qquad$ Rejection Region: $\qquad$

QA.22. A study compared three methods of making espresso: Bar Machine (BM, i=1), Hyper Espresso Method (HIP, $\mathrm{i}=2$ ), and I-Espressos System (IT, $\mathrm{i}=3$ ). There were $\mathrm{n}=9$ replicates per method ( $\mathrm{N}=27$ ). The following summary statistics were computed for the response Foam Index (\%). $\bar{y}_{1 \bullet}=32.4 \quad \bar{y}_{2 \bullet}=61.3 \quad \bar{y}_{3 \bullet}=39.7 \quad \mathrm{MSE}=71.53$ p.22.a. Use Tukey's method to compare all pairs of methods. BM IT HIP
p.22.b. Compute the minimum significant difference for all pairs of means based on the Bonferroni method.

QA.23. An experiment is conducted to compare $t=3$ diets for parrots. The diets are described as follow.
Diet 1: Corn Diet 2: Sunflower seeds Diet 3: Corn + Sunflower seeds
Give two orthogonal contrasts of interest among these 3 treatments (diets).

$$
l_{1}=\ldots \mu_{1}+\ldots \mu_{2}+\ldots \quad \mu_{3} \quad l_{2}=\ldots \quad \mu_{1}+\ldots \ldots \mu_{2}+\ldots \quad \mu_{3}
$$

QA.24. A Kruskal-Wallis test is conducted to compare 4 treatments, with $n=3$ replicates per treatment. The total of the 3 rank sums will be what?

QA.25. A study involved men's rating of attractiveness of women. A photograph of a woman was photoshopped so that the woman's t -shirt was one of 4 colors: White, Red, Blue, or Green. There were a total of $\mathrm{N}=120$ subjects, with subjects being randomly assigned to colors in a Completely Randomized ( $\mathrm{n}=30$ subjects per Treatment). The summary statistics are given below.
p.25.a. Complete the following ANOVA table. Is there evidence to conclude that color effects attractiveness ratings? Yes / No

| Color | n | Mean |
| :--- | ---: | ---: | ---: | ---: | ---: | :--- | :--- | :--- |
| White | $\mathbf{3 0}$ | 5.12 |
| Red | 30 | 5.95 |
| Blue | 30 | 5.07 |
| Green | 30 | 4.93 |
| Overall | 120 | 5.27 |$\quad |$| ANOVA |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
| Source | df | SS | MS | F |
| Trts (Color) |  |  |  |  |
| Error |  | 1.50 |  | \#N/A |
| Total |  |  | \#N/A |  |
|  |  |  | \#N/A | \#N/A |
| \#N/A |  |  |  |  |

p.25.b. Give a contrast comparing the Red Shirt mean with mean of the remaining 4 colors.

Contrast Coefficients: $l_{R v W B G}=$ $\qquad$ $\mu_{R} \longrightarrow \mu_{W}$ $\qquad$ $\mu_{B} \longrightarrow \mu_{G}$
p.25.c. Give the estimated contrast, its standard error, and the t -test for testing $H_{0}: l=0 \quad H_{A}: l \neq 0$
$\hat{l}=$ $\qquad$
$\hat{S E}\{\hat{l}\}=$ $\qquad$ Test Stat: $\qquad$ Rejection Region: $\qquad$

QA.26. A 1-Way ANOVA is fit with $\mathrm{t}=5$ treatments and $\mathrm{n}_{\mathrm{i}}=4$ replicates per treatment. The Mean Square Error is 300 . Compute Tukey's HSD and Bonferroni's Minimum Significant Difference for comparisons among all pairs of treatment means.

Tukey HSD = $\qquad$ Bonferroni MSD = $\qquad$

## Part B: Randomized Block Design

QB.1. A study is conducted to compare 4 varieties of cat food on weight gain in kittens. 4 Kittens are selected at random from each of 12 litters with 4 or more kittens. Of the 4 kittens selected from each litter, one is assigned to variety $A$, one to $B$, one to $C$, and one to $D$ (at random). Weight change at 16 weeks is obtained for each kitten. Complete the following ANOVA table and use Bonferroni's method to compare all pairs of variety (population) mean weight change.

| Source | df | SS | MS | F | Critical F |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Variety |  | 600 |  |  |  |
| Litter |  |  |  |  |  |
| Error |  | 990 |  |  |  |
| Total | 47 | 3000 |  |  |  |

Variety Means: A: 21
B: 28
C: 22
D: 27
$\mathrm{H}_{0}$ : No Variety Differences
$H_{A}$ : Variety Differences Exist
Test Statistic $\qquad$ Rejection Region $\qquad$
Critical t-value for Bonferroni's Method: $\qquad$
Standard error of Difference between 2 Variety Means:
$B_{i j}$
Comparison
Confidence Interval
Conclude
A vs B
A vs C
A vs D
$B$ vs $C$
$B$ vs D
C vs D
QB.2. An experiment is conducted to compare the effects of 4 types of fertilizer on the growth of a particular plant.
A sample of 8 locations (blocks) in a large yard are selected and 4 plants are planted at each location. At each location, the 4 plants are randomly assigned such that one receives fertilizer A, one receives fertilizer B, one receives fertilizer C, and one receives fertilizer D. Complete the following Analysis of Variance Table.

| Source | df | SS | MS | F | $\mathrm{F}(.05)$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Fertilizer |  | 395.8 |  |  |  |
| Location |  | 329.3 |  |  |  |
| Error |  |  |  |  |  |
| Total |  | 745.3 |  |  |  |

The means for the fertilizers are: $A=27.1, B=29.0, C=33.7, D=35.9$. Use Bonferroni's method to make pairwise comparisons among all pairs of varieties with an experimentwise error rate of 0.05

QB.3. A Randomized Block Design is conducted to compare the bioavailabilities of 4 formulations of a test drug. A sample of 8 subjects is obtained, and each subject receives each formulation once (in random order with adequate time between administrations of drug).

DF(Treatments) $\qquad$ DF(Block) $\qquad$ DF(Error) $\qquad$ DF(Total) $\qquad$
QB.4. A randomized block design is conducted to compare $t=3$ treatments in $b=4$ blocks. Your advisor gives you the following table of data form the experiment (she was nice enough to compute treatment, block, and overall means for you), where: $T S S=\sum(Y-\bar{Y})^{2}$

| Blk\Trt | 1 | 2 | 3 | BlkMean |
| :---: | :---: | :---: | :---: | :---: |
| 1 | 20 | 22 | 24 | 22 |
| 2 | 10 | 13 | 16 | 13 |
| 3 | 28 | 25 | 34 | 29 |
| 4 | 10 | 12 | 14 | 12 |
| TrtMean | 17 | 18 | 22 | 19 |
| TSS |  |  |  |  |
| 658 |  |  |  |  |

p.4.a. Complete the following ANOVA table:

| Source | df | SS | MS | F_obs | F(.05) | Reject HO: No Effect? |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Treatments |  |  |  |  |  |  |
| Blocks |  |  |  |  |  |  |
| Error |  |  |  |  |  |  |
| Total |  |  |  |  |  |  |

p.4.b. Compute the Relative Efficiency of having used a Randomized Block instead of a Completely Randomized Design $\operatorname{RE}(\mathrm{RB}, \mathrm{CR})=$ $\qquad$
p.4.c.. Compute Tukey's minimum significant difference for comparing all pairs of container types:

Tukey's $\mathrm{W}=$ $\qquad$
p.4.d. Give results graphically using lines to connect Trt Means that are not significantly different: T1 T2 T3

QB.5. Jack and Jill wish to compare the effects of 3 internet pop-up advertisements (ad1, ad2, ad3) on click throughs. Their response is the fraction of all website visitors who are exposed to the pop-up who click through (analyzed as clickthroughs per 1000 exposures). They identify a large number of potential websites that are comparable with respect to:complexity and traffic.
p.5.a. Jack conducts a Completely Randomized Design, sampling 60 websites and randomly assigns them so that 20 receive ad1, 20 receive ad 2 , and 20 receive ad 3 . He obtains the following results:
$\bar{y}_{1}=25 \quad \bar{y}_{2}=35 \quad \bar{y}_{3}=45 \quad S S_{T r t s}=4000 \quad S S_{E R R}=68400$
Give Jack's test for testing $\mathrm{H}_{0}$ : No advertisement effects:

## p.5.a.i. Test Statistic:

p.5.a.ii. Reject $\mathrm{H}_{0}$ if Jack's test statistic falls in the range $\qquad$
p.5.b. Jill conducts a Randomized Block Design, sampling 12 websites (blocks) and assigns each ad to each website (randomizing the order of the ads to the websites). She obtains the following results:
$\bar{y}_{1}=25 \quad \bar{y}_{2}=35 \quad \bar{y}_{3}=45 \quad S S_{\text {Trts }}=2400 \quad S S_{\text {Blocks }}=36000 \quad S S_{E R R}=11000$
Give Jill's test for testing $\mathrm{H}_{0}$ : No advertisement effects:
p.5.b.i. Test Statistic:
p.5.b.ii. Reject $H_{0}$ if Jill's test statistic falls in the range $\qquad$
p.5.c. Obtain Jack's and Jill's minimum significant differences based on Bonferroni's method for comparing all pairs of advertisement effects

Jack's $\mathrm{B}_{\mathrm{ij}}=$ $\qquad$ Jill's $B_{i j}=$ $\qquad$

QB.6. A study was conducted to compare 3 speed reduction marking (SRM) conditions on drivers' acceleration in an automobile simulator. A sample of 15 drivers was selected, and each driver drove the simulator under the 3 SRM conditions (No SRM, Longitudinal SRM, Traverse SRM).
p.6.a The following tables give the treatment (and overall) means, and the partial ANOVA table. Complete the ANOVA table and test $H_{0}: \mu_{N}=\mu_{\mathrm{L}}=\mu_{\mathrm{T}}$.

|  |  |
| :--- | ---: |
| Treatment | Mean |
| No SRM | 0.1613 |
| Longitudinal SRM | 0.1260 |
| Traverse SRM | -0.0320 |
| Overall | 0.0851 |$\quad$| ANOVA |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Source | df | SS | MS | F | F(.05) | Reject H0? |
| Trts |  |  |  |  |  |  |
| Drivers |  | 0.9865 |  |  |  |  |
| Error |  |  |  |  |  |  |
| Total |  | 2.3309 |  |  |  |  |

p.6.b Use Tukey's method to obtain simultaneous $95 \%$ confidence intervals for comparing all pairs of treatment means.

QB.7. An experiment was conducted to determine whether initiation times for cricket players are effected by ball color and illumination level. There were 6 treatments (combinations of ball color (Red/White) and Illumination level (571/1143/1714)). There were 5 subjects (blocks) who were observed under each condition. The mean initiation time for
each player under each condition (treatment) is given in the following table. Use Friedman's test to determine whether there are any significant differences among the treatment medians.

| Subject | Trt1 | Trt2 | Trt3 | Trt4 | Trt5 | Trt6 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 125 | 121 | 131 | 124 | 110 | 120 |
| 2 | 178 | 183 | 156 | 175 | 169 | 168 |
| 3 | 178 | 167 | 159 | 157 | 167 | 166 |
| 4 | 147 | 126 | 147 | 146 | 150 | 136 |
| 5 | 144 | 153 | 162 | 171 | 157 | 163 |

Test Statistic: $\qquad$ Rejection Region: $\qquad$

QB.8. An experiment was conducted to compare 4 brands of antiperspirant in terms of percentage sweat reduction. A sample of 24 subjects was obtained, and each subject was measured using each antiperspirant. Model:

$$
y_{i j}=\mu+\alpha_{i}+\beta_{j}+\varepsilon_{i j} \quad i=1, \ldots, 4 j=1, \ldots, 24 \quad \sum_{i=1}^{4} \alpha_{i}=0 \quad \beta_{j} \sim N\left(0, \sigma_{b}^{2}\right) \quad \varepsilon_{i j} \sim N\left(0, \sigma^{2}\right)
$$

p.8.a. The 4 antiperspirant brand mean y-values are given below. Compute the overall mean.

$$
\bar{y}_{1}=15.6 \quad \bar{y}_{2}=25.0 \quad \bar{y}_{3}=26.5 \quad \bar{y}_{4}=26.5 \quad \bar{y}=
$$

$\qquad$
p.8.b. Complete the following partial ANOVA table:

| ANOVA |  |  |  |  |  |
| :--- | :--- | ---: | :---: | :---: | :---: |
| Source | df | SS | MS | $F$ | $F(0.05)$ |
| Subject |  | 14183.5 |  | \#N/A | \#N/A |
| Brand |  | 1976.75 |  |  |  |
| Error |  | 11740.25 |  | \#N/A | \#N/A |
| Total |  | 27900.5 | \#N/A | \#N/A | \#N/A |

p.8.c. Test $H_{0}$ : No differences among Brand Effects $\quad H_{A}$ : Differences exist among brands
p.8.c.i. Test Stat:___ p.2.c.ii. Reject $\mathrm{H}_{0}$ if Test $\operatorname{Stat}$ is in the range__ p.2.c.iii. P-value $>$ or $<.05$ ?
p.8.d. Use Tukey's Honest Significant Difference method to determine which (if any) brand means are significantly different.

Tukey's $\mathrm{W}=$ $\qquad$
p.8.e. Compute the Relative efficiency of the Randomized Block Design (relative to Completely Randomized Design). How many subjects would be needed per treatment (in CRD) to have the same standard errors of sample means as RBD.

Relative Efficiency = $\qquad$ \# of subjects per treatment in CRD $\qquad$

QB.9. A study was conducted to compare total distance covered by soccer players over a 16 minute game on fields of various sizes. The field sizes were $30 \times 20$ meters, $40 \times 30$, and $50 \times 40$. A sample of 8 skilled soccer players were selected and are treated as blocks for this analysis. The total distance covered by the 8 players on the 3 field sizes are given in the following table. Use Friedman's test to test whether true mean distance covered differs among the 3 field sizes.

| Player | $30 \times 20$ | $40 \times 30$ | $50 \times 40$ |
| :--- | :--- | :--- | :--- |
| 1 | 1141 | 1558 | 1493 |
| 2 | 1573 | 1963 | 2036 |
| 3 | 1802 | 2140 | 2218 |
| 4 | 1745 | 2142 | 2078 |
| 5 | 1663 | 2116 | 2036 |
| 6 | 1288 | 1748 | 1696 |
| 7 | 1705 | 2105 | 2167 |
| 8 | 1340 | 1755 | 1748 |

Friedman's Test Statistic $\qquad$ Rejection Region: $\qquad$ P-value < or > . 05

QB.10. A study compared $t=4$ warm-up protocols in terms of vertical jump ability in dancers. There were $b=10$ dancers, each dancer was measured under each warm-up protocol and the experiment is a Randomized Block Design with dancers as blocks.
$y_{i j}=\mu+\alpha_{i}+\beta_{j}+\varepsilon_{i j}=\mu_{i}+\beta_{j}+\varepsilon_{i j}$
The treatments and their means are: Static Stretch: 38.0 Dynamic Stretch: 41.4 Static\&Dynamic Stretch: 41.0 Control: 37.8
p.10.a. Complete the following ANOVA table.

| Source | df | SS | MS | F_obs | F(.05) |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Treatment (Warm-up) |  |  |  |  |  |
| Block (Dancer) |  | 850.6 |  |  |  |
| Error |  | 42.0 |  |  |  |
| Total |  |  |  |  |  |

p.10.b. Do you reject $H_{0}: \alpha_{1}=\ldots=\alpha_{4}=0 \quad\left(\mu_{1}=\ldots=\mu_{4}\right) ? \quad$ Yes $\quad / \quad$ No
p.10.c. Compute the Relative Efficiency of the RCB to the Completely Randomized Design. How many subjects would be needed per treatment to have the same standard error of a treatment (warm-up protocol) in a CRD?

Relative Efficiency $\qquad$ \# of Subjects per treatment $\qquad$
p.10.d. Compute Bonferroni's minimum significant difference and determine which treatments are significantly different.

$$
\begin{array}{llll}
\text { Control } & \text { Static } & \text { Static\&Dynamic } & \text { Dynamic }
\end{array}
$$

QB.11. An experiment was conducted comparing various treatments (involving various hydrocolloids and amounts of wheat flower) with the goal of reducing oil content in a food product. The experiment was conducted in separate replicates (blocks). One response measured was Oil Content of the sample. The partial ANOVA table is given below.

| Source | df | SS | MS | F | F(.05) |
| :--- | ---: | ---: | :--- | :--- | :---: |
| Treatments | 12 | 261.146 |  |  |  |
| Blocks | 2 | 0.523 |  | \#N/A | \#N/A |
| Error |  | 0.689 |  | \#N/A | \#N/A |
| Total |  | 262.358 | \#N/A | \#N/A | \#N/A |

p.11.a. Complete the table. Is the P-value for testing $\mathrm{H}_{0}$ : No Treatment Effect $>0.05$ or $<0.05$
p.11.b. Give the number of Treatments and number of Blocks in the experiment. \# Trts $=$ $\qquad$ \# Blks = $\qquad$
p.11.c. What is the estimated standard error of the difference between any 2 treatment means? $\quad S E\left\{\bar{Y}_{i \bullet}-\bar{Y}_{j \bullet}\right\}$
p.11.d. Suppose we wish to use Scheffe's method to compare all pairs of treatment means. What would be the minimum significant difference?

QB.12. An experiment was conducted as a Randomized Block Design with 3 treatments (Weight Belts: None (Control), Air Belt, and Comp Vest Belt) in 12 blocks (Subjects). The response was the maximum acceptable work load. The mean squares for Treatments (Belts), Blocks (Subjects), and Error (Trt/Block Interaction) and the Belt means are given below.

Treatments: $\mathrm{MST}=131.3$ Blocks: $\mathrm{MSB}=3628.9$ Error: $\operatorname{MSE}=38.2 \quad \bar{Y}_{1 \bullet}=34.45 \quad \bar{Y}_{2 \bullet}=38.84 \quad \bar{Y}_{3 \bullet}=40.93$ p.12.a. Use Tukey's method to compare the all pairs of belt means.

Tukey's HSD: $\qquad$ No Belt Air Belt Comp Vest Belt
p.12.b. Compute the Relative Efficiency of the Randomized Block Design to the Completely Randomized Design.
$R E=$ $\qquad$
p.12.c. How many subjects would be needed per treatment in a Completely Randomized Design to have the same precision in terms of the difference between mean that was obtained in this experiment? How many total?

Subjects per Treatment $\qquad$ Total Subjects $\qquad$

## Part C: Latin Square Design

QC.1. An experiment was conducted to compare 5 treatments (Seed Rate) in a latin square design. A field was partitioned into 5 rows and 5 columns, such that each treatment appeared in each row once, and each column once. The response is grain yield.

|  | level | rowmean | colmean | trtmean |
| ---: | ---: | ---: | ---: | ---: |
|  | 1 | 54.15 | 52.43 | 47.13 |
|  | 2 | 56.30 | 54.30 | 51.72 |
|  | 3 | 52.29 | 54.44 | 55.73 |
| 4 | 52.58 | 55.30 | 59.17 |  |
|  | 5 | 57.31 | 56.16 | 58.88 |


| ANOVA |  |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- | :---: |
| Source | df | SS | MS | F | F(0.05) |
| Seed Rate |  | 522.74 |  |  |  |
| Field Row |  | 99.13 |  | \#N/A | \#N/A |
| Field Column |  | 38.60 |  | \#N/A | \#N/A |
| Error |  |  |  | \#N/A | \#N/A |
| Total |  | 716.61 | \#N/A | \#N/A | \#N/A |

p.1.a. Complete the ANOVA table.
p.1.b. Test $H_{0}$ : No differences among Seed Rate Effects $\quad H_{A}$ : Differences exist among Seed rates
p.1.b.i. Test Stat:
p.3.b.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.3.c.iii. P-value > or < . 05 ?
p.1.c. Use Bonferroni's method to determine which (if any) Seed Rates are significantly different.

Bonferroni's B = $\qquad$
p.1.d. Compute the Relative efficiency of the Latin Square Design (relative to Completely Randomized Design). Relative Efficiency = $\qquad$

QC.2. A latin square design is conducted comparing sales of juice in 5 container types (Treatment factor). The experiment is conducted in 5 stores (row blocking factor), over 5 weeks (column blocking factor) in a manner such that each container is sold in each store once, and each week once. Results of sales are given below:

Container Means: C1: 80
C2: 100
C3: 90
C4: 60
C5: 85
$\mathrm{SS}_{\text {Row }}=1000$
$\mathrm{SS}_{\text {Column }}=400 \mathrm{SS}_{\text {Error }}=240$
p.2.a. Compute the Relative Efficiency of having used a Latin Square instead of a Completely Randomized Design
$\operatorname{RE}(\mathrm{LS}, \mathrm{CR})=$ $\qquad$
p.2.b. Compute Bonferroni’s minimum significant difference for comparing all pairs of container types:

Bonferroni's $\mathrm{B}=$ $\qquad$
p.2.c. Give results graphically using lines to connect Containers that are not significantly different:

C 4 C 1 C 5 C 3 C 2

QC3. Researchers conducting a Latin Square Design with $t=5$ treatments, row blocks, and column blocks report a relative efficiency (relative to completely randomized design) of 3 . How many replicates per treatment would they need if they conducted this experiment as a completely randomized design to have equivalently precise standard errors of sample means as they obtained from the latin square?

QC.4. A study was conducted as a latin square design to compare countries of origin (Treatments: 1=Australia, 2=Chile, 3=France, 4=New Zealand) on ratings of quality of wine. There were 120 raters ( 30 per Week (Row Factor)) and 4 Bottle Labels (Column Factor)). The response for each "cell" in the design was the sum of the quality ratings of the 30 subjects for that week/label /country.
p.4.a Complete the following ANOVA table (hint: there are 16 total measurements).

| ANOVA |  |  |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- | :--- | :--- |
| Source | df | SS | MS | F | F(.05) | Reject H0? |
| Week (Row) |  | 729.2 |  |  |  |  |
| Label (Column) |  | 414.7 |  |  |  |  |
| Country (Trt) |  | 1937.7 |  |  |  |  |
| Error |  | 397.9 |  |  |  |  |
| Total |  | 3479.4 |  |  |  |  |

p.4.b The sample means for the 4 countries are: Australia=89.5, Chile=116.5, France=94.25, New Zealand=90.5. Use Tukey's method to compare all pairs of country means.

QC.5. A latin square design was used to test for treatment effects among 5 mixes of concrete in terms of tensile strenth. There were 5 molds, and 5 workers who made and poured the concrete molds. The design is shown below.

|  | Worker1 | Worker2 | Worker3 | Worker4 | Worker5 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Mold1 | Mix1 | Mix2 | Mix3 | Mix4 | Mix5 |
| Mold2 | Mix2 | Mix3 | Mix4 | Mix5 | Mix1 |
| Mold3 | Mix3 | Mix4 | Mix5 | Mix1 | Mix2 |
| Mold4 | Mix4 | Mix5 | Mix1 | Mix2 | Mix3 |
| Mold5 | Mix5 | Mix1 | Mix2 | Mix3 | Mix4 |

p.5.a. The mix means are: $70,65,50,80$, and 75 for Mixes $1-5$, respectively. The error sum of squares is $S S E=600$. Use Bonferroni's method to compare all pairs of Mix means.

Bonferoni's MSD: $\qquad$ Mix3 Mix2 Mix1 Mix5 Mix4
p.5.b. The sums of squares for molds and workers are $S S R=1000$ and $S S C=400$, respectively. Compute the Relative Efficiency of the Latin Square design, relative to the Completely Randomized Design.
$\operatorname{RE}(\mathrm{LS}, \mathrm{CR})=$ $\qquad$

QC.6. A researcher is using a latin square design to compare 4 brands of car tires in terms of miles driven before reaching a given wear level. For one blocking factor they use tire position (Driver Front, Passenger Front, Driver Rear, Passenger Rear). They choose to use 12 cars as the other blocking factor. Note that each brand will be on each car once, and on each tire position 3 times. What will be the error degrees of freedom?

QC.7. A Latin Square Design is used for an experiment with 3 treatments. The column (blocking) factor has 3 levels and the row blocking factor has 12 levels (thus, multiple squares have been formed). Give the critical F-value for testing for treatment effects.
$\mathrm{F}_{95 ; \mathrm{dfl} 1 \mathrm{dr} 2}=$ $\qquad$

QC.8. A marketing experiment was conducted as a latin square with $t=6$ treatments (Shelf Space, with levels $2,4,6,8,10,12$ feet) conducted in $t=6$ Rows (Stores) over $t=6$ Columns (Weeks). Each shelf space was run in each store one time and during each week one time. The response was number of packages of baking soda sold during the week in the store.
p.8.a. Complete the partial ANOVA table is given below. Is $\mathrm{H}_{0}$ : No Difference in Shelp Space Effects Rejected? Yes / No

| ANOVA |  |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- | :---: |
| Source | df | SS | MS | F | F(.05) |
| Store |  | 8287 |  | \#N/A | \#N/A |
| Week |  | 982 |  | \#N/A | \#N/A |
| ShelfSpace |  | 395 |  |  |  |
| Error |  | 1488 |  | \#N/A | \#N/A |
| Total |  | 11151 | \#N/A | \#N/A | $\# N / A$ |

p.8.b. Compute the minimum significant difference for comparing Shelf Space Means, based on Bonferroni’s method. The smallest and largest means were 40.5 for 2 feet, and 51.2 for 10 feet. Are they significantly different?

Minimum Significant Difference $\qquad$ 2 and 10 feet significantly different? Yes / No

## Part D: 2-Factor (and Higher) Crossed ANOVA

QD.1. Based on the 2014 WNBA season, we have the point totals ( Y ) by game Location (Home/Away) for a sample of 10 Players. Each player played 17 home games and 17 away games. Consider the model:
$y_{i j k}=\mu+\alpha_{i}+\beta_{j}+\alpha \beta_{i j}+\varepsilon_{i j k} \quad i=1, \ldots, a \quad j=1, \ldots, b \quad k=1, \ldots, r \quad \sum_{i=1}^{a} \alpha_{i}=0 \quad \beta_{j} \sim N\left(0, \sigma_{b}^{2}\right) \quad \alpha \beta_{i j} \sim N\left(0, \sigma_{a b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$

| ANOVA |  |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- | :--- |
| Source | df | SS | MS | F | F(0.05) |
| Player |  | 3879.30 |  |  |  |
| Home |  | 1.30 |  |  |  |
| P*H |  | 323.67 |  |  |  |
| Error |  | 15787.29 |  | \#N/A | \#N/A |
| Total |  | 19991.56 | \#N/A | \#N/A | \#N/A |

$$
E\{M S E\}=\sigma^{2} \quad E\{M S A B\}=\sigma^{2}+r \sigma_{a b}^{2} \quad E\{M S B\}=\sigma^{2}+r \sigma_{a b}^{2}+a r \sigma_{b}^{2} \quad E\{M S A\}=\sigma^{2}+r \sigma_{a b}^{2}+\frac{b r \sum_{i=1}^{a} \alpha_{i}^{2}}{a-1}
$$

p.1.a. Complete the partial ANOVA table.
p.1.b. Test whether there is an interaction between Player and Location (Home). $\mathrm{H}_{0}: \sigma_{\mathrm{ab}}{ }^{2}=0$
p.1.b.i. Test Stat: $\qquad$ p.4.b.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.4.b.iii. P-value >or < .05?
p.1.c. Test whether there is Location (Home vs Away) Main Effect. $H_{0}: \alpha_{1}=\alpha_{2}=0$
p.1.c.i. Test Stat: $\qquad$ p.4.c.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.4.c.iii. P-value > or < . 05 ?
p.1.d. Test whether there is Player Main Effect. $\mathrm{H}_{0}: \sigma_{b}{ }^{2}=0$
p.1.d.i. Test Stat: $\qquad$ p.4.d.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.4.d.iii. P-value >or < . 05 ?
p.1.e. Give unbiased estimates of each of the variance components:
$\qquad$

$$
\hat{\sigma_{b}^{2}}=
$$

$\qquad$ $\hat{\sigma^{2}}=$ $\qquad$

QD. 2 The broiler chicken study had 60 replicates at each of 2 levels of factor A (Base: Sorghum or Corn) and 2 levels of Factor B (Methionine: Present or Absent). One response reported was the weight of the wing drumette.

Model: $y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad i=1,2 j=1,2 k=1, \ldots, 60 \quad \sum_{i=1}^{2} \alpha_{i}=\sum_{j=1}^{2} \beta_{j}=\sum_{i=1}^{2}(\alpha \beta)_{i j}=\sum_{j=1}^{2}(\alpha \beta)_{i j}=0 \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.2.a. The following table gives the means (SDs) for each treatment:

| Base\Meth | Absent | Present | Mean |
| :---: | :---: | :---: | :---: |
| Sorghum | $46.4(8.0)$ | $34.8(6.0)$ | 40.6 |
| Corn | $38.8(6.0)$ | $41.6(10.0)$ | 40.2 |
| Mean | 42.6 | 38.2 | 40.4 |

Complete the following ANOVA table:

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Base |  |  |  |  |  |
| Methionine |  |  |  |  |  |
| B*M |  | 3110.4 |  |  |  |
| Error |  | 13924 |  | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A | \#N/A |

p.2.b. Test $\mathrm{H}_{0}$ : No Interaction between Base and Methionine
p.2.b.i. Test Stat: $\qquad$ p.2.b.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.2.b.iii. P-value > or < .05?
p.2.c. Test $\mathrm{H}_{0}$ : No Base effect
p.2.c.i. Test Stat: $\qquad$ p.2.c.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.2.c.iii. P-value >or < .05?
p.2.d. Test $\mathrm{H}_{0}$ : No Methionine effect
p.2.d.i. Test Stat: $\qquad$ p.1.d.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.2.d.iii. P-value >or < .05?

QD.3. A study is conducted to compare 5 methods of oiling bowling alleys (Factor A) on scores by professional bowlers. A random sample of 10 professional bowlers (Factor B) are observed twice on each of these 5 oiling methods (the scores are totals pins over 7 games/100). These are the only oiling methods of interest. The partial ANOVA table is given below for the model:

$$
Y_{i j k}=\mu+\alpha_{i}+b_{j}+(a b)_{i j}+\varepsilon_{i j k} \sum \alpha_{i}=0 \quad b_{j} \sim N\left(0, \sigma_{b}^{2}\right) \quad(a b)_{i j} \sim N\left(0, \sigma_{a b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)
$$

| Source | df | SS | MS | F | Critical F |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Oiling Method | 4 | 43.0 | 10.75 |  |  |
| Bowler | 9 | 10.1 | 1.12 |  |  |
| OilxBowler | 36 | 24.3 | 0.675 |  |  |
| Error | 50 | 37.5 | 0.75 |  |  |
| Total | 99 | 114.9 |  |  |  |

Conduct the following tests:
p.3.a $H_{0}$ : No bowler/oiling method interaction: $\sigma_{a b}{ }^{2}=0 \quad H_{A}$ : bowler/oiling method interaction: $\sigma_{a b}{ }^{2}>0$

Test Statistic $\qquad$ Rejection Region $\qquad$ Do you conclude there is a significant interaction? $\qquad$
p.3.b $H_{0}$ : No Oiling Method Differences: $\alpha_{1}=\alpha_{2}=\alpha_{3}=\alpha_{4}=\alpha_{5}=0 \quad H_{A}$ : Differences exist among oil methods (Not all $\alpha_{i}=0$ ) Test Statistic $\qquad$ Rejection Region $\qquad$ conclude there is a significant oiling method effect? $\qquad$
p.3.c. $H_{0}$ : No bowler effect: $\sigma_{b}{ }^{2}=0 \quad H_{A}$ : bowler effect exists: $\sigma_{b}{ }^{2}>0$

Test Statistic $\qquad$ Rejection Region $\qquad$ Conclude there is a significant interaction? $\qquad$

QD.4. An experiment is conducted to measure the effects of 4 weave types and 3 test speeds on the breaking strength of fibers. Four replicates are obtained at each combination of weave type and test speed. These are the only weave types and fibers of interest to the researchers.
p.4.a Complete the following ANOVA table, and conduct the tests for interactions and main effects.

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Weave Type |  | 3224.82 |  |  |  |
| Test Speed |  | 3186.53 |  |  |  |
| Interaction |  | 20.98 |  |  |  |
| Error |  | 389.28 |  |  |  |
| Total |  | 6821.62 |  |  |  |

p.4.b $H_{0}$ : No Interaction between weave type and test speed $\quad$ Reject $H_{0} \quad / \quad$ Fail to Reject $H_{0}$
p.4.c $H_{0}$ : No weave type effects
p.4.d. $\mathrm{H}_{0}$ : No test speed effects

Reject $\mathrm{H}_{0}$ / Fail to Reject $\mathrm{H}_{0}$

Reject $\mathrm{H}_{0} \quad$ Fail to Reject $\mathrm{H}_{0}$

QD.5. An experiment is conducted to determine the effects of 3 ripening stages (Factor A) and 2 screw speeds (Factor B) on Water Solubility Index in Bananas. There are 3 replicates at each combination of levels of Factors A and B. The Sample means are given in the following table, as well as row and column means, and the partial ANOVA table. Both factors are considered fixed in this design.
Means

| Factor <br> A\B | 1 | 2 | RowMean |
| :---: | :---: | :---: | :---: |
| 1 | 23.4 | 24.1 | 23.75 |
| 2 | 24.3 | 25.8 | 25.05 |
| 3 | 25.4 | 26.2 | 25.80 |
| CoIMean | 24.37 | 25.37 | 24.87 |

ANOVA

| Source | df | SS | MS | F_obs | F(.05) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| A |  | 12.91 |  |  |  |
| B |  |  |  |  |  |
| A*B |  | 0.57 |  |  |  |
| Error |  | 1.53 |  |  |  |
| Total |  |  |  |  |  |

p.5.a. Complete the ANOVA Table.
p.5.b. Test $\mathrm{H}_{0}$ : No Interaction between ripening stage (A) and screw speed (B).
p.5.b.i. Test Stat: $\qquad$ p.5.b.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.5.b.iii. P-value >or < . 05 ?
p.5.c. Compute Tukey's minimum significant difference (W) and Bonferroni's minimum significant difference (B) when we wish to compare all 3 ripening stages at a given level of screw speed. (That is, when screw speed=1, or when screw speed $=2$ ).

Tukey's W: $\qquad$ Bonferroni's B $\qquad$

QD.6. An experiment was conducted to measure variability in gauge readings among operators (Factor A) and parts from a production process (Factor B). The 3 operators are the only ones at the company, so they are fixed. The 20 parts are a random sample from many parts produced, so they are random. Each operator makes $r=2$ measurements per part.
p.6.a. Assuming the mixed model with fixed operator effects and random (and independent) parts and interaction effects, complete the following ANOVA table:

| Source | df | SS | MS | F_obs | F(.05) | Reject H0: No Effect? |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Operators |  | 2.6 |  |  |  |  |
| Parts |  | 1185.4 |  |  | 1.867 |  |
| O*P |  | 27.1 |  |  | 1.603 |  |
| Error |  |  |  |  |  |  |
| Total |  | 1274.6 |  |  |  |  |

p.6.b. Assuming no Operator/Part interaction, based on Tukey's method, how far apart would 2 operators means need to differ by to be considered significantly different, when we simultaneously compare all pairs of operators? Tukey's $\mathrm{W}=$ $\qquad$
QD. 7 Based on the following Analysis of Variance table, based on a balanced 2-Way ANOVA, answer the following questions.

| ANOVA |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Source | df | SS | MS | F_obs | P-value |
| A | 4 | 600 | 150 | 7.5 | 0.0001 |
| B | 3 | 270 | 90 | 4.5 | 0.0065 |
| A*B | 12 | 360 | 30 | 1.5 | 0.1495 |
| Error | 60 | 1200 | 20 |  |  |
| Total | 79 | 2430 |  |  |  |

p.7.a. Number of levels of Factor A $\qquad$
p.7.b. Number of levels of Factor B $\qquad$
p.7.c. Number of replicates per treatment (combination of levels of A and B) $\qquad$
p.7.d. Estimate of standard deviation of measurements within same treatment $\qquad$
p.7.e. P-value for test of $\mathrm{H}_{0}$ : No interaction between the effects of levels of factors A and B $\qquad$
p.7.f. P-value for test of $\mathrm{H}_{0}$ : No effects among levels of factor B $\qquad$
p.7.g. Number of pairs of levels of factor A in a multiple comparison procedure $\qquad$
QD. 8 Consider the following table from a 2-Factor Fixed Effects Model

Table 4. Ash content (g kg-1 DM) of selected saltgrass accessions grown during 10 weeks in water culture at four salinity levels

Salinity leve1 (dS m-1)
$\begin{array}{lllll}\text { Accession } & 1.5 & 10 & 30 & 50\end{array}$

| AL1 | $65 \pm 2.2$ | $78 \pm 4.1$ | $90 \pm 2.9$ | $98 \pm 3.1$ |
| :--- | :--- | :--- | :--- | :--- |
| AL3 | $63 \pm 2.8$ | $79 \pm 5.0$ | $92 \pm 4.3$ | $94 \pm 4.9$ |
| Arg1 | $84 \pm 3.1$ | $99 \pm 4.4$ | $102 \pm 3.7$ | $107 \pm 3.1$ |
| Arg2 | $86 \pm 3.6$ | $94 \pm 3.5$ | $96 \pm 3.2$ | $102 \pm 3.8$ |
| CA1 | $72 \pm 2.4$ | $88 \pm 2.9$ | $103 \pm 3.5$ | $105 \pm 3.7$ |
| CA4 | $66 \pm 2.5$ | $84 \pm 4.2$ | $89 \pm 4.1$ | $89 \pm 4.7$ |
| CA13 | $70 \pm 1.9$ | $90 \pm 3.3$ | $97 \pm 3.6$ | $96 \pm 4.0$ |
| CA17 | $68 \pm 2.6$ | $85 \pm 4.4$ | $94 \pm 4.0$ | $94 \pm 3.8$ |
| CH1 | $79 \pm 3.4$ | $94 \pm 3.6$ | $103 \pm 4.4$ | $106 \pm 5.6$ |
| CH2 | $75 \pm 3.1$ | $95 \pm 4.8$ | $100 \pm 5.9$ | $99 \pm 4.1$ |
| CT2 | $71 \pm 2.9$ | $84 \pm 3.1$ | $90 \pm 3.7$ | $92 \pm 4.2$ |
| DE1 | $75 \pm 1.8$ | $88 \pm 2.9$ | $96 \pm 3.5$ | $99 \pm 3.3$ |
| DE3 | $74 \pm 2.5$ | $82 \pm 2.6$ | $90 \pm 3.0$ | $91 \pm 3.4$ |
| GA2 | $56 \pm 1.4$ | $77 \pm 1.9$ | $86 \pm 2.2$ | $88 \pm 2.9$ |
| GA3 | $62 \pm 1.6$ | $79 \pm 2.4$ | $89 \pm 2.8$ | $90 \pm 3.2$ |
| GA6 | $66 \pm 1.9$ | $76 \pm 1.8$ | $81 \pm 2.7$ | $82 \pm 3.5$ |
| a | Values are means $\pm$ SE of six rep1icates |  |  |  |
| p.8.a. Give the degrees of freedom for the Analysis of Variance. |  |  |  |  |


| Source | df |
| :--- | :---: |
| Accession |  |
| Salinity |  |
| A*S |  |
| Error |  |
| Total |  |

p.8.b. Set up the calculation of the Error Sum of Squares (SE represents standard error of the mean)

QD.9. An experiment is conducted to compare 4 varieties of cheddar cheese (fixed effect) in terms of bitterness scores among a sample of 5 raters (random effect). Each rater tastes each variety twice (they are not told which variety they are tasting, and bitterness is rated on a visual analogue scale, ranging from 0 (low) to 10 (high).).

Model:

$$
y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad i=1, \ldots, 4 j=1, \ldots, 5 k=1, \ldots, 2 \quad \sum_{i=1}^{4} \alpha_{i}=0 \quad \beta_{j} \sim N\left(0, \sigma_{b}^{2}\right) \quad(\alpha \beta)_{i j} \sim N\left(0, \sigma_{a b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)
$$

p.9.a. The sample means for the 4 varieties are: $\bar{y}_{A}=4.0 \quad \bar{y}_{B}=7.0 \quad \bar{y}_{C}=3.0 \quad \bar{y}_{D}=6.0$

Complete the following ANOVA table:

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Variety |  |  |  |  |  |
| Rater |  | 24.01 |  |  |  |
| V*R |  | 3.83 |  |  |  |
| Error |  |  |  | \#N/A | \#N/A |
| Total |  | 152.05 | \#N/A | \#N/A | \#N/A |

p.9.b. Test $\mathrm{H}_{0}$ : No Interaction between variety and rater
p.9.b.i. Test Stat: $\qquad$ p.9.b.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.9.b.iii. P-value >or < . 05 ?
p.9.c. Compute Tukey's minimum significant difference (W) and Bonferroni's minimum significant difference (B) when we wish to test for differences among variety effects.

Tukey's W: $\qquad$ Bonferroni's B $\qquad$

QD.10. A research paper reported the following Analysis of Variance, based on raw data presented earlier in the paper. There were 3 factors ( $A, B, C$ ), each at 2 levels. There were 3 blocks, each block being made up of one observation for each of the $2(2)(2)=8$ combinations of factors A, B, and C. Note that the Block and Corrected Total Sums of Squares are correct in the table. Ignore $\mathbf{B C}$ and Residual for this table.

| Response 3 |  | ncy |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| anova for selec | factorial n |  |  |  |  |  |
| Analysis of variance | le [Partial | of sq | - Type |  |  |  |
|  | Sum of |  | Mean | F | p-value |  |
| Source | Squares | df | Square | Value | Prob $>$ F |  |
| Block | 7.07 | 2 | 3.53 |  |  |  |
| Model | 19.37 | 4 | 4.84 | 11.69 | $\leqslant 0.0001$ | significant |
| A-Charge Voitage | 3.25 | 1 | 3.25 | 7.85 | 0.0123 |  |
| B-Armature Mass | 8.94 | 1 | 8.94 | 21.60 | 0.0002 |  |
| C-Starting Location | 5.25 | 1 | 5.25 | 12.69 | 0.0024 |  |
| BC | 1.92 | 1 | 1.92 | 4.64 | 0.0459 |  |
| Residual | 7.04 | 17 | 0.41 |  |  |  |
| Cor Total | 33.47 | 23 |  |  |  |  |

q.10.a. Suppose we want to obtain the sums of squares for each of the 3 main effects (recall that each has 2 levels). How many observations are each of these means based on? That is, the mean of all observations when factor $A$ is at its low level is based on how many observations (This will be the same as $A_{\text {High, }} \mathrm{B}_{\text {Low, }}, \mathrm{B}_{\text {High, }}, \mathrm{C}_{\text {Low }}, \mathrm{C}_{\text {High }}$ )?
q.10.b. The means for each of the levels for $A, B$, and $C$ are given below, as well as the overall mean. Compute the sums of squares for $\mathrm{A}, \mathrm{B}$, and C , respectively. How do they compare to those given by the authors (who were probably using more decimal places internally).

| Group | Mean |
| :--- | ---: |
| A_Low | 5.03 |
| A_High | 4.89 |
|  |  |
| B_Low | 5.57 |
| B_High | 4.35 |
|  |  |
| C_Low | 5.43 |
| C_High | 4.49 |
|  |  |
| Overall | 4.96 |

Actual SSA = $\qquad$ SSB $=$ $\qquad$ SSC = $\qquad$

QD.11. An experiment was conducted to determine the effects of viewing a magical film versus a non-magical film in children. Samples of 326 -year olds, and 328 -year-olds were selected, and randomly assigned such that 16 of each agegroup viewed the magical film and 16 of each age-group viewed the non-magical film (both age and film category are treated as fixed effects). The following table gives the means (SDs) for each treatment. The response (y) was a score on an imagination scale (rating of a child acting out an object or animal).

Model: $y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad i=1,2 j=1,2 k=1, \ldots, 16 \quad \sum_{i=1}^{2} \alpha_{i}=\sum_{j=1}^{2} \beta_{j}=\sum_{i=1}^{2}(\alpha \beta)_{i j}=\sum_{j=1}^{2}(\alpha \beta)_{i j}=0 \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.11.a. The following table gives the means (SDs) for each treatment:

| Age\Film | Non-Magical | Magical | Mean |
| :---: | :---: | :---: | :---: |
| 6 | $17.0(2.7)$ | $21.6(4.1)$ | 19.3 |
| 8 | $18.7(3.8)$ | $22.7(3.8)$ | 20.7 |
| Mean | 17.85 | 22.15 | 20 |

Complete the following ANOVA table:

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Age |  |  |  |  |  |
| Film |  |  |  |  |  |
| A*F |  | 1.44 |  |  |  |
| Error |  | 794.7 |  | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A | \#N/A |

p.11.b.i. Test Stat: $\qquad$ p.11.b.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.11.b.iii. P-value > or < .05?
p.11.c. Test $\mathrm{H}_{0}$ : No Age effect
p.11.c.i. Test Stat: $\qquad$ p.11.c.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.11.c.iii. P-value > or < . 05 ?
p.11.d. Test $\mathrm{H}_{0}$ : No Film Type effect
p.11.d.i. Test Stat: $\qquad$ p.11.d.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.11.d.iii. P-value > or < .05?

QD.12. A study was conducted, measuring the effects of 3 electronic Readers and 4 Illumination levels on time for people to read a given text ( 100 s of seconds). There were a total of 60 subjects, 5 each assigned to each combination of Reader/Illumination level. For this analysis, consider both Reader and Illumination level as fixed effects.
p.12.a Complete the following ANOVA table, and test for significant Reader/Illumination Interaction effects, as well as main effects for Reader and Illumination levels.

| ANOVA |  |  |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- | :--- | :--- |
| Source | df | SS | MS | F | F(0.05) | Significant Effects? |
| Reader |  | 70.70 |  |  |  |  |
| Illumination |  | 148.11 |  |  |  |  |
| Read*Illum |  | 2.15 |  |  |  |  |
| Error |  | 365.02 |  |  |  |  |
| Total |  | 585.98 |  |  |  |  |

p.12.b Use Tukey's Method to make all pairwise comparisons among Readers.
p.12.c Use Bonferroni's method to make all pairwise comparisons among Illumination levels.

QD.13. An experiment had 2 factors, each with 2 levels: Factor A:Instructional Method (Standard and Enhanced), and Factor B: Instructional Medium (Desktop and Mobile Device). There were a total of $\mathrm{N}=88$ subjects, with $\mathrm{n}=22$ receiving each treatment (combination of method and medium). The sample means and standard deviations of scores on a transfer test are given below.

| Mean | Medium |  |  |  | SD | Medium |  |
| :--- | ---: | ---: | ---: | ---: | :--- | :--- | ---: |
| Method | Desktop | Mobile | Overall |  | Method | Desktop | Mobile |
| Standard | 2.58 | 2.36 | 2.47 |  | Standard | 1.93 | 1.76 |
| Enhanced | 4.04 | 4.34 | 4.19 |  | Enhanced | 3.04 | 2.67 |
| Overall | 3.31 | 3.35 | 3.33 |  |  |  |  |

The model fit is a 2-Way fixed effects ANOVA with interaction. $y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k}$
p.13.a. Compute the sums of squares for: Method, Medium, and complete the following ANOVA table.

| Source | df | SS | MS | F_obs | F(0.05) |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Method |  |  |  |  |  |
| Medium |  |  |  |  |  |
| $\mathrm{M}^{*}$ M |  | 1.4872 |  |  |  |
| Error |  | 487.053 |  |  |  |
| Total |  |  |  |  |  |

p.13.b. Do you reject the hypothesis: $H_{0}^{A B}:(\alpha \beta)_{11}=(\alpha \beta)_{12}=(\alpha \beta)_{21}=(\alpha \beta)_{22}=0 \quad$ Yes $\quad / \quad$ No
p.13.c. Do you reject the hypothesis: $H_{0}^{A}: \alpha_{1}=\alpha_{2}=0 \quad$ Yes $/$ No
p.13.d. Do you reject the hypothesis: $H_{0}^{B}: \beta_{1}=\beta_{2}=0 \quad$ Yes $/$ No

QD.14. Compute the appropriate minimum significant difference for comparing (fixed) treatment means.
Two-way ANOVA: Factor A Fixed, $a=4$, Factor B Random, $b=5, n=3$ replicates per combination of factors A and B.
$S S A=600 \quad S S B=1000 \quad S S A B=600 \quad S S E=1600$
Compute Tukey's HSD for comparing levels of factor A.

QD.15. An experiment was conducted to compare $a=3$ theories for the apparent modulus of elasticity ( $Y$ ) of $b=3$ apple varieties. The 3 theories were: Hooke's, Hertz's, and Boussineq's; the 3 apple varieties were: Golden Delicious, Red Delicious, and Granny Smith. The researchers determined the elasticity for $r=15$ apples based on each combination of theory and variety. For the purposes of this experiment, each factor is fixed.

$$
\begin{aligned}
& \text { Model: } y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad \varepsilon_{i j k} \sim \operatorname{NID}\left(0, \sigma^{2}\right) \quad \sum_{i=1}^{a} \alpha_{i}=\sum_{j=1}^{b} \beta_{j}=\sum_{i=1}^{a}(\alpha \beta)_{i j}=\sum_{j=1}^{b}(\alpha \beta)_{i j}=0 \\
& \sum_{i=1}^{a} \sum_{j=1}^{b} \sum_{k=1}^{n}\left(y_{i j k}-\bar{y}_{i j \bullet}\right)^{2}=17.095 \sum_{i=1}^{a} \sum_{j=1}^{b} \sum_{k=1}^{n}\left(y_{i j k}-\bar{y}_{\ldots . .}\right)^{2}=113.119 \quad \text { bn } \sum_{i=1}^{a}\left(\bar{y}_{i \bullet \bullet}-\bar{y}_{\ldots .}\right)^{2}=57.987 \quad \text { an } \sum_{j=1}^{b}\left(\bar{y}_{\bullet j \bullet}-\bar{y}_{\ldots .}\right)^{2}=35.779
\end{aligned}
$$

| Cell Means | GoldenDelicious | RedDelicious | GrannySmith | Row Mean |
| :--- | :---: | :---: | :---: | :---: |
| Hooke | 2.68 | 3.46 | 4.23 | 3.457 |
| Hertz | 2.44 | 3.06 | 3.84 | 3.113 |
| Boussinesq | 1.53 | 1.89 | 2.36 | 1.927 |
| Column Mean | 2.217 | 2.803 | 3.477 | 2.832 |

Complete the following Analysis of Variance Table, and test for interaction effects and main effects.

| Source | df | SS | MS | F | F(.95) | P-value |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Theory |  |  |  |  |  | $>0.05$ or <0.05 |
| Variety |  |  |  |  |  | $>0.05$ or $<0.05$ |
| Theory*Variety |  |  |  |  |  | $>0.05$ or $<0.05$ |
| Error |  |  |  | \#N/A | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A | \#N/A | \#N/A |

QD.16. A 2-Way Random Effects model is fit, where a sample of $a=8$ products were measured by a sample of $b=6$ machinists, with $r=3$ replicates per machinist per product. The model fit is as follows (independent random effects):
$Y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad \alpha_{i} \sim \operatorname{NID}\left(0, \sigma_{a}^{2}\right) \quad \beta_{j} \sim \operatorname{NID}\left(0, \sigma_{b}^{2}\right) \quad(\alpha \beta)_{i j} \sim \operatorname{NID}\left(0, \sigma_{a b}^{2}\right) \quad \varepsilon_{i j k} \sim \operatorname{NID}\left(0, \sigma^{2}\right)$

You are given the following sums of squares: $\quad S S A=420 \quad S S B=350 \quad S S A B=140 \quad S S E=210$

Give the test statistic and rejection region for the following 3 tests. Note for test 1, your rejection region will be symbolic, give the specific numerator and denominator degrees of freedom. Also give unbiased (ANOVA) estimates of each variance component.

1) $H_{0}^{A B}: \sigma_{a b}^{2}=0 \quad H_{A}^{A B}: \sigma_{a b}^{2}>0$
2) $H_{0}^{A}: \sigma_{a}^{2}=0$
$H_{A}^{A}: \sigma_{a}^{2}>0$
3) $H_{0}^{B}: \sigma_{b}^{2}=0 \quad H_{A}^{B}: \sigma_{b}^{2}>0$

1: Test Stat: $\qquad$ Rejection Region: $\qquad$ Estimate: $\qquad$
2: Test Stat: $\qquad$ Rejection Region: $\qquad$ Estimate: $\qquad$
3: Test Stat: $\qquad$ Rejection Region: $\qquad$ Estimate: $\qquad$

QD.17. A 2-Way (crossed) ANOVA is used to measure the effects of 2 factors, each at 3 levels. There are 4 replicates for each treatment (combination of levels of factors A and B). There is a significant interaction between the 2 factors, so the researchers choose to use Tukey's method to compare all pairs of treatment means. Give Tukey's W for comparing all pairs of means, with MSE $=100$.

QD.18. An unbalanced two-way ANOVA was conducted to compare desire scores for a product (Y) among a sample of female consumers. The factors were Product ( $\mathrm{X}_{1}=1$ if Shoe, -1 if Ink toner) and Endorser ( $\mathrm{X}_{2}=1$ if Celebrity, -1 if Online Consumee). The sample sizes were: $S / C=43, S / O=44, I / C=46, I / O=43$. Four regressions models were fit:

Model 1: $E\left\{Y_{i j k}\right\}=\beta_{0}+\beta_{1} X_{1 i j k}+\beta_{2} X_{2 i j k}+\beta_{3} X_{1 i j k} X_{2 i j k} \quad S_{1}=169.1 \quad \hat{Y}_{1}=2.478-0.183 X_{1}+0.088 X_{2}+0.148 X_{1} X_{2}$
Model 2: $E\left\{Y_{i j k}\right\}=\beta_{0}+\beta_{1} X_{1 i j k}+\beta_{2} X_{2 i j k}$
$S S E_{2}=172.9 \quad Y_{2}=2.472-0.181 X_{1}+0.086 X_{2}$
Model 3: $E\left\{Y_{i j k}\right\}=\beta_{0}+\beta_{2} X_{2 i j k}+\beta_{3} X_{1 i j k} X_{2 i j k}$
$S S E_{3}=174.9 \quad \hat{Y}_{3}=2.479+0.092 X_{2}+0.146 X_{1} X_{2}$
p.18.a. Test whether there is an interaction between Product and Endorser. $\mathrm{H}_{0}$ : $\qquad$ $\mathrm{H}_{\mathrm{A}}$ : $\qquad$

Test Statistic $\qquad$ Rejection Region $\qquad$ P -value $>\mathbf{0 . 0 5}$ or $<\mathbf{0 . 0 5}$
p.18.b. Test whether there is main effect for Product. $\mathrm{H}_{0}$ : $\qquad$ $\mathrm{H}_{\mathrm{A}}: \longrightarrow$
$\qquad$

Test Statistic $\qquad$ Rejection Region $\qquad$ P-value $>\mathbf{0 . 0 5}$ or $<\mathbf{0 . 0 5}$
p.18.c. Based on Model 1, give the predicted scores for all combinations of Product and Endorser.


QD.19. An experiment was conducted to compare rating scores assigned to wines. There were 5 wine producers (vineyards), there were 2 appelations (white and red wines), and 14 judges. Each judge rated each wine variety a single time. Note that each producer makes a white wine and a red wine. Give the sources of variation and their corresponding degrees of freedom, using the highest order interaction as the error term.

Source
Degrees of Freedom

QD.20. An experiment is conducted as a 2-Way factorial design with $a=3$ levels, $b=2$ levels, and $n=4$ replicates per treatment. $M S E=100$. There is a significant interaction, so the researchers decide to use Tukey's method to compare all of the combinations of levels of Factors A and B. Compute Tukey's HSD for the researchers.
$\operatorname{HSD}=$ $\qquad$

QD.21. A study was conducted to compare 3 methods of measuring blood pressure (Factor A, Fixed). There were 20 Human Subjects in the study (Factor B, Random). Each subject was measured twice by each method ( $n=2$ ).

Model: $Y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad \sum_{i=1}^{a} \alpha_{i}=0 \quad \beta_{j} \sim N\left(0, \sigma_{\beta}^{2}\right) \quad(\alpha \beta)_{i j} \sim N\left(0, \sigma_{\alpha \beta}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.21.a. Complete the ANOVA table.

| ANOVA |  |  |  |
| :--- | :--- | ---: | :--- |
| Source | df | SS | MS |
| Method |  | 4117 |  |
| Subject |  | 20120 |  |
| Meth*Subj |  | 1680 |  |
| Error |  | 1666 |  |
| Total |  | 27583 | \#N/A |

p.21.b. Test $H_{0}^{A B}: \sigma_{\alpha \beta}^{2}=0$ (No interaction between methods and subjects) vs $H_{A}^{A B}: \sigma_{\alpha \beta}^{2}>0$

Note: $\mathrm{F}\left(.05, \mathrm{df}_{1}, \mathrm{df}_{2}\right)=1.603$
Test Statistic: $\qquad$ Rejection Region $\qquad$
p.21.c. Test $H_{0}^{A}: \alpha_{1}=\alpha_{2}=\alpha_{3}=0 \quad\left(\right.$ No method effect differences) vs $H_{A}^{A}:$ Not all $\alpha_{i}=0$

Test Statistic: $\qquad$ Rejection Region $\qquad$
p.21.d. Test $H_{0}^{B}: \sigma_{\beta}^{2}=0$ (No subject effect differences) vs $H_{A}^{B}: \sigma_{\beta}^{2}>0$

Test Statistic: $\qquad$ Rejection Region $\qquad$

QD.22. An experiment was conducted to measure the energy efficiency of electric clothes dryer machines. Factor A was Clothing Type ( $1=$ Towels, $2=$ Jeans, $3=$ Thermal Clothing) and Factor B was Dryer Type ( $1=$ Electric Dryer, $2=\mathrm{Bi}-$ directional Electric dryer, $3=$ Town Gas-Fired Dryer, $4=$ LPG-Fired dryer). There were $n=3$ replicates per combination of Clothing and Dryer Types. The model is given below, with Factors A and B both being Fixed factors.
$Y_{i j k}=\mu+\alpha_{i}+\beta_{j}+\alpha \beta_{i j}+\varepsilon_{i j k} \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right) \quad \sum_{i} \alpha_{i}=\sum_{j} \beta_{j}=\sum_{i} \alpha \beta_{i j}=\sum_{j} \alpha \beta_{i j}=0$

| Source | sumsq | df | MS | F | F(.05) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Clothing Type |  |  |  |  |  |
| Dryer Type | 5.2296 |  |  |  |  |
| CxD | 0.0362 |  |  |  |  |
| Error |  |  |  | \#N/A | \#N/A |
| Total | 5.6011 |  | \#N/A | \#N/A | \#N/A |
|  |  |  |  |  |  |
|  | i | 1 | 2 | 3 | All |
|  | Mean | 1.5408 | 1.7118 | 1.6141 | 1.6222 |
|  | j | 1 | 2 | 3 | 4 |
|  | Mean | 1.3031 | 1.3437 | 2.2250 | 1.5871 |

p.22.a. Complete the Analysis of Variance Table.
p.22.b. For each of the following three null hypotheses, are they rejected?
$H_{0}^{A B}: \alpha \beta_{11}=\ldots=\alpha \beta_{34}=0$
$H_{0}^{A}: \alpha_{1}=\ldots=\alpha_{3}=0$ $\qquad$ $H_{0}^{B}: \beta_{1}=\ldots=\beta_{4}=0$ $\qquad$
p.22.c. Compute Tukey's HSD for comparing the all pairs among the 3 clothing types (use MSE and $\mathrm{df}_{\mathrm{E}}$ directly from the ANOVA table). Which pairs (if any) are significantly different?
p.22.d. Compute Bonferroni's MSD for comparing the all pairs among the 4 dryer types (use MSE and $\mathrm{df}_{\mathrm{E}}$ directly from the ANOVA table). Which pairs (if any) are significantly different?

QD.23. An unbalanced 2-Factor study was conducted to test for gender and movie type effects on brand recall of products placed in movies. Factor A was gender with $\mathrm{a}=2$ levels (female/male) and Factor B was movie type with $\mathrm{b}=3$ levels (action/comedy/drama). The response was Y (number of correct recalls of items, $0-6$ ). The model was fit as a linear regression and set up as follows. There were a total of $\mathrm{n}=137$ subjects in the study.
$X_{1}=\left\{\begin{array}{c}1 \text { if Female } \\ -1 \text { if Male }\end{array} \quad X_{2}=\left\{\begin{array}{c}1 \text { if Action } \\ -1 \text { if Drama } \\ 0 \text { if Comedy }\end{array} \quad X_{3}=\left\{\begin{array}{c}1 \text { if Comedy } \\ -1 \text { if Drama } \\ 0 \text { if Action }\end{array}\right.\right.\right.$
Model 1 contains all main effects and interactions, Model 2 contains all main effects, Models 3 and 4 contain only Factor $A$ and $B$ main effects, respectively.

Model 1: $E\{Y\}=\mu+\alpha_{1} X_{1}+\beta_{1} X_{2}+\beta_{2} X_{3}+\alpha \beta_{11} X_{1} X_{2}+\alpha \beta_{12} X_{1} X_{3} \quad S S E_{1}=231.66 \quad d f_{E 1}=131$
Model 2: $E\{Y\}=\mu+\alpha_{1} X_{1}+\beta_{1} X_{2}+\beta_{2} X_{3} \quad S S E_{2}=232.04 \quad d f_{E 2}=133$
Model 3: $E\{Y\}=\mu+\alpha_{1} X_{1} \quad S S E_{3}=243.62 \quad d f_{E 3}=135 \quad \operatorname{Model} 4: E\{Y\}=\mu+\beta_{1} X_{2}+\beta_{2} X_{3} \quad S S E_{4}=236.44 \quad d f_{E 4}=134$
p.23.a. Use Models 1 and 2 to show that the gender/movie type interaction is not significant. $H_{0}^{A B}: \alpha \beta_{11}=\alpha \beta_{12}=0$

Test Statistic $\qquad$ Rejection Region $\qquad$ $\mathrm{P}>$ or $<.05$
p.23.b. Use Models 2, 3, and 4 to test whether gender and/or movie type main effects are significant.
$H_{0}^{A}: \alpha_{1}=0 \quad H_{0}^{B}: \beta_{1}=\beta_{2}=0$

Factor A: Test Statistic $\qquad$ Rejection Region $\qquad$ $\mathrm{P}>$ or $<.05$

Factor B: Test Statistic $\qquad$ Rejection Region $\qquad$ $\mathrm{P}>$ or $<.05$

QD.24. In the Broiler Chicken study, Factor A is base diet with $\mathrm{a}=2$ levels (sorghum and corn), and Factor B is methionine with $b=2$ levels (absent and present). There were $n=60$ chickens per treatment in a Completely Randomized Design. The model fit is given below, along with treatment means. The Error sum of squares is $S S E=167560$.

$$
Y_{i j k}=\mu+\alpha_{i}+\beta_{j}+\alpha \beta_{i j}+\varepsilon_{i j k} \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right) \quad \sum_{i} \alpha_{i}=\sum_{j} \beta_{j}=\sum_{i} \alpha \beta_{i j}=\sum_{j} \alpha \beta_{i j}=0
$$

|  | Meth- | Meth+ | RowMean |
| :--- | ---: | ---: | ---: |
| Sorghum | 46 | 35 |  |
| Corn | 39 | 42 |  |
| ColMean |  |  |  |

p.24.a. Compute SSA, SSB, and SSAB, and complete the following Analysis of Variance Table.

| Source | df | SS | MS | F | F(.05) | Significant? |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| A |  |  |  |  |  | Yes / No |
| B |  |  |  |  |  | Yes / No |
| AB |  |  |  |  |  | Yes / No |
| Error |  |  |  | \#N/A | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A | \#N/A | \#N/A |

p.24.b. Assuming the interaction effect is significant, compute Tukey's HSD and Bonferroni's MSD among the 4 treatments. (Hint: for Bonferroni's method there are $4(4-1) / 2=6$ pairs of treatments. Based on Tukey's HSD identify which pairs of treatments are not significantly different by joining lines or letter superscripts.

Tukey HSD $\qquad$ Bonferroni MSD $\qquad$ Sorg/Meth+ Corn/Meth- Corn/Meth+ Sorg/Meth-

QD.25. A wine making experiment was conducted to determine the effects of two factors on Color index (Y, chroma) of Chardonnay Wine. Factor A was pressing method (aerobic/inert) and factor B was handling condition (oxidative/reductive). There were $\mathrm{n}=3$ replicates for each combination of factors A and B . The treatment means are given below (Note that this would be considered a Fixed Effects model).

$$
Y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad \sum_{i} \alpha_{i}=\sum_{j} \beta_{j}=\sum_{i}(\alpha \beta)_{i j}=\sum_{j}(\alpha \beta)_{i j}=0 \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)
$$

| Mean |  |  |  |
| :--- | ---: | ---: | ---: |
| A $\backslash \mathrm{B}$ | Oxidative | Reductive | Row Mean |
| Aerobic | 5.88 | 6.37 | 6.125 |
| Inert | 6.97 | 7.46 | 7.215 |
| Column Mean | 6.425 | 6.915 | 6.67 |

p.25.a. Complete the following ANOVA table.

| Source | df | SS | MS | F | F(.05) |
| :--- | :--- | :--- | :--- | :--- | :--- |
| A |  |  |  |  |  |
| B |  |  |  |  |  |
| AB |  |  |  |  |  |
| Error |  | 1.2456 |  | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A | \#N/A |

p.25.b. What are your conclusions based on the following tests?
$H_{0}^{A B}:(\alpha \beta)_{11}=(\alpha \beta)_{a b}=0$ $\qquad$
$H_{0}^{A}: \alpha_{1}=\ldots=\alpha_{a}=0$ $\qquad$ $H_{0}^{B}: \beta_{1}=\ldots=\beta_{b}=0$ $\qquad$

QD.26. An experiment was conducted to study the effects of maturation time (Factor A: 1,2,3 months) and type of storage (Factor B: Glass, Plastic with Light Toast chips, Plastic with Medium Toast Chips, Plastic with Dark Toast Chips, Oak Barrel). The model is fit as an additive 2-factor, crossed Fixed Effects Model. The response was Bitterness (Y, in IBUs) and there were $n=3$ replicates per treatment.
$Y_{i j k}=\mu+\alpha_{i}+\beta_{j}+\varepsilon_{i j k} \quad \sum_{i} \alpha_{i}=\sum_{j} \beta_{j}=0 \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
The marginal means and Error Sum of Squares are given below.
$\operatorname{SSE}=16.6 \quad$ Maturation Time: $\bar{y}_{1 .}=18.1 \quad \bar{y}_{2 . .}=20.1 \quad \bar{y}_{3 . .}=18.9$
Storage: $\bar{y}_{.10}=18.5 \quad \bar{y}_{.2 .}=19.4 \quad \bar{y}_{.3 .}=19.0 \quad \bar{y}_{.4 \cdot}=19.9 \quad \bar{y}_{.5 \cdot}=18.3$
p.26.a.Use Tukey's method to compare all pairs of Maturation Times.

1 Month 3 Months 2 Months
p.26.b. Use Bonferroni's to compare all pairs of Storage Types.
B
MT
LT
DT

QD.27. A mixed model was fit relating measured blood alcohol content ( $\mathrm{Y}=100^{*} \mathrm{~g} / 210 \mathrm{~L}$ ) among Factor A ( $\mathrm{a}=6$ Models of Instruments, Fixed factor) and Factor B ( $b=3$ Subjects who consumed alcohol, Random factor), with $\mathrm{n}=10$ measurements made by each instrument on each subject. The model fit is as follows (independent random effects):

$$
Y_{i j k}=\mu+\alpha_{i}+\beta_{j}+(\alpha \beta)_{i j}+\varepsilon_{i j k} \quad \sum_{i=1}^{a} \alpha_{i}=0 \quad \beta_{j} \sim N I D\left(0, \sigma_{b}^{2}\right) \quad(\alpha \beta)_{i j} \sim N I D\left(0, \sigma_{a b}^{2}\right) \quad \varepsilon_{i j k} \sim N I D\left(0, \sigma^{2}\right)
$$

p.27.a. Complete the following ANOVA table and test for main effects and interactions.
$H_{0}^{A B}: \sigma_{a b}^{2}=0 \quad H_{0}^{A}: \alpha_{1}=\ldots=\alpha_{a}=0 \quad H_{0}^{B}: \sigma_{b}^{2}=0$

| Source | df | SS | MS | F | $\mathrm{F}(.05)$ | P-value |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Instrument (A) |  | 20.58 |  |  |  | $>$ |
| Subject (B) |  | 210.5 |  |  |  | $>$ |
| Inst x Subj (AB) |  | 3.52 |  |  |  | $>$ |
| Error |  | 8.89 |  | \#N/A | \#N/A | .05 |
| Total |  | 243.49 | \#N/A | \#N/A | \#N/A | \#N/A |

p.27.b. Obtain point estimates for $\sigma^{2}, \sigma_{a b}^{2}, \sigma_{b}^{2}$
$s^{2}=$ $\qquad$ $s_{a b}^{2}=$ $\qquad$ $s_{b}^{2}=$ $\qquad$
p.27.c. Compute Tukey's HSD (W) for comparing all pairs of instruments.

Tukey's HSD (W) = $\qquad$

QD.28. An unbalanced experiment was conducted to study the effects of Selenium ( $\mathrm{X}_{1}=1$ if Present, 0 if Absent) and N -acetyl-cysteine ( $\mathrm{X}_{2}=1$ if Present, 0 if Absent) on sperm count in infertile men. The response was sperm count (Y). The sample sizes in the 4 cells were slightly unbalanced due to the nature of the study, with a total sample size of $N=420$. The following 4 models are fit.

Model 1: $E\{Y\}=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\beta_{12} X_{1} X_{2} \quad Y=60.80+5.20 X_{1}+3.95 X_{2}+1.25 X_{1} X_{2} \quad S S E_{1}=98074 \quad R_{1}^{2}=.3985$
Model 2: $E\{Y\}=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2} \quad \hat{Y}=60.80+5.19 X_{1}+3.94 X_{2} \quad \operatorname{SSE}=98730 \quad R_{2}^{2}=.3914$
Model 3: $E\{Y\}=\beta_{0}+\beta_{1} X_{1} \quad \hat{Y}=60.78+5.19 X_{1} \quad S S E_{2}=105263 \quad R_{3}^{2}=.3117$
Model 4: $E\{Y\}=\beta_{0}+\beta_{2} X_{2} \quad \hat{Y}=60.78+3.94 X_{2} \quad S S E_{2}=110061 \quad R_{4}^{2}=.2367$
p.28.a. Test whether there is an interaction between Selenium and N -acetyl-cystine on sperm count.
$\mathrm{H}_{0}$ : $\qquad$ Test Stat: $\qquad$ Rej Region: $\qquad$ $\mathrm{P}>$ or $<.05$
p.28.b. Given the additive model (Model 2), test whether there is a Selenium effect.
$\mathrm{H}_{0}$ : $\qquad$ Test Stat: $\qquad$ Rej Region: $\qquad$ $\mathrm{P}>$ or $<.05$
p.28.c. Given the additive model (Model 2), test whether there is a N -acetyl-cystine effect.
$\mathrm{H}_{0}$ : $\qquad$ Test Stat: $\qquad$ Rej Region: $\qquad$ $\mathrm{P}>$ or $<.05$

## Part E: Nested Designs

QE.1. A study is conducted to compare 3 types of traffic signal settings (pre-timed, semi-actuated, and fully actuated). A sample of 30 intersections in a large city are obtained, and 10 are assigned to each of the 3 settings at random. Measurements are obtained at each signal at 20 "points" in time, where $Y=$ traffic delay (seconds/vehicle). Write out the sources of variation and degrees of freedom for the ANOVA table. Would these factors each be best described as fixed or random? What would be the appropriate error term for testing for signal effects? What would be the degrees of freedom?

QE.2. A study is conducted to compare pH levels in rivers in 3 geographic areas. Random samples of 5 rivers were sele within each of the geographic areas, and 4 replicates were obtained within each river.
p.2.a. Complete the following Analysis of Variance table.

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Area |  | 4000 |  |  |  |
| River w/in Area |  | 2400 |  |  |  |
| Error |  | 2250 |  |  |  |
| Total |  |  |  |  |  |

p.2.b. Compute Bonferroni's B to be used to compare all pairs of geographic areas.

QE.3. An experiment was conducted to compare 3 traffic light types (Factor A). A random sample of 9 intersections (Factor B) were selected, and 3 were assigned to each traffic light type at random. Types are treated as fixed, and intersections are to be treated as random. Measurements of average waiting times are made at each intersection over $\mathbf{r}=$ 8 time periods. Set up the ANOVA table, giving all sources of variation, degrees of freedom, F-statistics (symbolically by specifying appropriate Mean squares), and critical F-values.

Source | df |
| :---: |
| F_obs $=$ MS1/MS2 |$\quad$ F(.05)

QE.4. An experiment was conducted to compare 5 machines in terms of strain-readings (y) of glass-cathode supports. The engineer had 4 "heads" from which the glass was formed for each machine (that is, the 4 "heads" for machine 1 differ from those from machine 2, etc..., implying "heads" are nested under machines). Each "head" is measured 4 times (replicates) to obtain a strain reading. Note that these are the only 5 machines of interest (fixed effects), but the "heads" used are a sample from a larger population of "heads" (random effects).

Model: $y_{i j k}=\mu+\alpha_{i}+\beta_{j(i)}+\varepsilon_{i j k} \quad i=1, \ldots, 5 j=1, \ldots, 4 k=1, \ldots, 4 \quad \sum_{i=1}^{5} \alpha_{i}=0 \quad \beta_{j(i)} \sim N\left(0, \sigma_{b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.4.a. The 5 machine total strain-readings are: $93,81,82,88$, and 58 , respectively. Compute the machine and overall means (hint: how many measurements are taken from each machine):

$$
\bar{y}_{1}=\ldots \quad \bar{y}_{2}=\ldots \quad \bar{y}_{3}=\quad \bar{y}_{4}=\quad \bar{y}_{5}=\quad \bar{y}=
$$

p.4.b. Complete the following partial ANOVA table:

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Machine |  |  |  |  |  |
| Head(M) |  | 282.88 |  |  |  |
| Error |  | 642 |  | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A | \#N/A |

p.4.c.. Test $\mathrm{H}_{0}$ : differences among "head" effects $\left(\sigma_{b}^{2}=0\right)$ vs $\mathrm{H}_{\mathrm{A}}$ : Differences among "head" effects $\left(\sigma_{b}^{2}>0\right)$ p.4.c.i. Test Stat: $\qquad$ p.4.c.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.4.c.iii. P-value >or < .05?
p.4.c. Compute Tukey's minimum significant difference (W) and Bonferroni's minimum significant difference (B) when we wish to test for differences among machine effects.

Tukey's W: $\qquad$ Bonferroni's B $\qquad$

QE.5. An experiment was conducted to compare 6 batches of auto body side panels in terms of deviations from nominal position (y). The engineer samples 2 "groups" of body panels from each batch (that is, the 2 "groups" for batch 1 differ from those from batch 2, etc..., implying "groups" are nested under batches). Each "group" has 3 individual body panels selected and measured (replicates) for $y$. Note that these are a random sample of batches (random effects), and the "groups" used are a sample from a larger population of "groups" (random effects).

Model: $y_{i j k}=\mu+\alpha_{i}+\beta_{j(i)}+\varepsilon_{i j k} \quad i=1, \ldots, 6 j=1,2 k=1,2,3 \quad \alpha_{i} \sim N\left(0, \sigma_{a}^{2}\right) \quad \beta_{j(i)} \sim N\left(0, \sigma_{b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.5.a. The 6 batch mean $y$-values are given below. Compute the overall mean, and obtain the sum of squares for batches.
$\begin{array}{llllll}\bar{y}_{1}=4.000 & \bar{y}_{2}=2.017 & \bar{y}_{3}=-4.567 & \bar{y}_{4}=-1.117 & \bar{y}_{5}=4.050 & \bar{y}_{6}=-1.100\end{array} \quad \bar{y}=$ $\qquad$
p.5.b. Complete the following partial ANOVA table:

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Batch |  |  |  |  |  |
| Grp(B) |  | 62.05 |  |  |  |
| Error |  | 438.57 |  | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A | \#N/A |

p.5.c. Test $\mathrm{H}_{0}$ : differences among "group" effects $\left(\sigma_{b}^{2}=0\right)$ vs $\mathrm{H}_{\mathrm{A}}$ : Differences among "group" effects $\left(\sigma_{b}^{2}>0\right)$
p.5.c.i. Test Stat: $\qquad$ p.5.c.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.5.c.iii. P-value > or < . 05 ? p.5.d. Test $\mathrm{H}_{0}$ : differences among "batch" effects $\left(\sigma_{a}^{2}=0\right)$ vs $\mathrm{H}_{\mathrm{A}}$ : Differences among "batch" effects $\left(\sigma_{a}^{2}>0\right)$ p.5.d.i. Test Stat: $\qquad$ p.5.d.ii. Reject $\mathrm{H}_{0}$ if Test Stat is in the range $\qquad$ p.5.d.iii. P-value > or < . 05 ?
p.5.e. The Expected Mean Squares for Batches, Groups within Batches, and Error are:
$E\{$ MSBatches $\}=\sigma^{2}+3 \sigma_{b}^{2}+3(2) \sigma_{a}^{2} \quad E\{\operatorname{MSGrp}(B)\}=\sigma^{2}+3 \sigma_{b}^{2} \quad E\{M S E\}=\sigma^{2}$
Give unbiased estimates of each of the variance components:

$$
\hat{\sigma_{a}^{2}}=\ldots \hat{\sigma_{b}^{2}}=\ldots \quad \hat{\sigma^{2}}=
$$

QE.6. Researchers conducted an experiment measuring acoustic metric values in 3 habitats ( $1=$ Cliff, $2=\mathrm{Mud}, 3=$ Gravel) in 3 patches, nested within each habitat, with replicates representing 5 sites within each patch ( $\mathrm{N}=3(3)(5)=45$ ). The habitats are considered to be fixed levels, while patches within habitats are considered to be random. The response measured was snap amplitude.
p.6.a Complete the following ANOVA table, and test for habitat effects ( $H_{0}: \alpha_{1}=\alpha_{2}=\alpha_{3}=0$ ) and for patch effects ( $H_{0}$ : $\sigma_{a b}{ }^{2}=0$ ).

| ANOVA |  |  |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- | :--- | :--- |
| Source | df | SS | MS | F | F(.05) | Reject H0? |
| Habitat |  | 403.3 |  |  |  |  |
| Patch(Hab) |  | 304.7 |  |  |  |  |
| Error |  | 386.6 |  |  |  |  |
| Total |  | 1094.6 |  |  |  |  |

p.6.b Compute Bonferroni's minimum significant difference for comparing pairs of habitat means.
p.6.c Obtain point estimate for $\sigma_{\mathrm{ab}}{ }^{2}$ and $\sigma^{2}$

QE.7. An engineering experiment was conducted to measure variation in semiconductors for a particular measurement. A sample of 5 lots (batches) of semiconductors was selected. Within each lot, 2 wafers were sampled. Each wafer was measured at 9 random sites (these are replicates). Note that the wafers are nested within lots. Both lots and wafers are random effects.
$Y_{i j k}=\mu+\alpha_{i}+\beta_{j(i)}+\varepsilon_{i j k} \quad \alpha_{i} \sim N\left(0, \sigma_{a}^{2}\right) \quad \beta_{j(i)} \sim N\left(0, \sigma_{b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$

| Source | df | SS | MS | F_obs | F(.05) |
| :--- | ---: | ---: | :--- | :--- | :--- |
| Lot |  | 1698.5 |  |  |  |
| Wafer(Lot) |  | 272.2 |  |  |  |
| Error |  | 1803.8 |  |  |  |
| Total |  |  |  |  |  |

p.7.b. Do you reject the hypothesis: $H_{0}^{A}: \sigma_{a}^{2}=0$ ? Yes / No
p.7.c. Do you reject the hypothesis: $H_{0}^{B}: \sigma_{b}^{2}=0$ ? Yes / No
p.7.d. For the nested design, with random factors A , and $\mathrm{B}(\mathrm{A})$, we have:

$$
E\{M S E\}=\sigma^{2} \quad E\{\operatorname{MSB}(A)\}=\sigma^{2}+n \sigma_{b}^{2} \quad E\{M S A\}=\sigma^{2}+n \sigma_{b}^{2}+b n \sigma_{a}^{2}
$$

Obtain unbiased estimates for the 3 variances:
$\hat{\sigma}=$ $\qquad$
$\qquad$
^ 2
$\sigma_{a}=$ $\qquad$

QE.8. For the following scenario compute the appropriate minimum significant difference for comparing (fixed) treatment means:
Nested 2-Way ANOVA: Factor A Fixed, $\mathrm{a}=3$, Factor B Fixed, $\mathrm{b}=3$, $\mathrm{n}=5$ replicates cell.
$S S A=120 \quad S S B(A)=180 \quad S S E=144$
Compute Bonferroni's MSD for comparing levels of factor A.

QE.9. A wildlife researcher is interested in comparing levels of a chemical in the water among the 4 lakes in a state park. The lakes are broken into many subsections based on a survey. She samples 3 subsections from each lake at random, and takes water measurements at 8 sites within each subsection. A laboratory measures the chemical in each of the water specimens. The lake means are: $76,72,60$, and 64 , respectively. The model is:
$Y_{i j k}=\mu+\alpha_{i}+\beta_{j(i)}+\varepsilon_{i j k} \quad \sum_{i=1}^{a} \alpha_{i}=0 \quad \beta_{j(i)} \sim N\left(0, \sigma_{b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.9.a. Compute the sum of squares, degrees of freedom, and mean square for lakes (factor A ).

SSA = $\qquad$ $\mathrm{df}_{\mathrm{A}}=$ $\qquad$ MSA = $\qquad$
The sum of squares for subsections (factor $B$ ) nested within lakes (factor $A$ ) is $\operatorname{SSB}(A)=640$, and the error sum of squares is $\mathrm{SSE}=2520$.
p.9.b. Test $H_{0}: \alpha_{1}=\ldots=\alpha_{4}=0$

Test Statistic: $\qquad$
$\qquad$ P -value $<$ or > . 05
p.9.c. Test $H_{0}: \sigma_{b}^{2}=0 \quad H_{A}: \sigma_{b}^{2}>0$

Test Statistic: $\qquad$ Rejection Region: $\qquad$ P-value < or > . 05

QE.10. An experiment was conducted to compare 2 methods of constructing blue jeans (Manually and with Laser Beams). Samples of 20 pairs of jeans were constructed by each method, and 3 measurements were made on each pair of jeans. Note that the blue jeans (random) are nested within the method (fixed) by which they were constructed. The statistical model is as follows (the response is the extension of the blue jeans).
$Y_{i j k}=\mu+\alpha_{i}+\beta_{j(i)}+\varepsilon_{i j k} \quad i=1,2 ; j=1, \ldots, 20 ; k=1,2,3 \quad \alpha_{1}+\alpha_{2}=0 \quad \beta_{j(i)} \sim N\left(0, \sigma_{\beta}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.10.a. Complete the following ANOVA table.

| ANOVA |  |  |  |
| :--- | :--- | ---: | :--- |
| Source | df | SS | MS |
| Method |  | 1185 |  |
| Jeans(Method) |  | 5851 |  |
| Error |  | 1837 |  |
| Total |  | 8873 | \#N/A |

p.10.b. Test whether there is a difference in the two methods. $H_{0}: \alpha_{1}=\alpha_{2}=0$

Test Statistic $\qquad$ Rejection Region $\qquad$ p -value $>0.05$ or $<0.05$
p.10.c. Obtain point estimates for $\sigma_{\beta}^{2}$ and $\sigma^{2}$
$\wedge^{2}$
$\sigma_{\beta}=$ $\qquad$ $\sigma^{2}=$ $\qquad$
p.10.d. The sample means for the two methods are $\bar{Y}_{1 \bullet \bullet}=65.1$ and $\bar{Y}_{2 \bullet}=71.3$. Compute a $95 \%$ Confidence Interval for $\alpha_{1}-\alpha_{2}$

95\% CI: $\qquad$

QE.11. An experiment was conducted to determine the effect of Twitter use (based on academic and co-curricular discussions). The experiment consisted of 8 class sections, 4 sections were placed in the Experimental condition (Twitter based course activities) and the remaining 4 sections were placed in the Control condition (no course based Twitter activities). There were $\mathrm{n}=18$ students (replicates) in each section. Note that the Twitter (Experimental/Control) factor would be considered Fixed, while the Class Sections are Random. The student's course grade on 4 point scale is Y.

The model is: $Y_{i j k}=\mu+\alpha_{i}+\beta_{j(i)}+\varepsilon_{i j k} \quad \sum_{i=1}^{a} \alpha_{i}=0 \quad \beta_{j(i)} \sim N\left(0, \sigma_{b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.11.a. Compute the sum of squares, degrees of freedom, and mean square for Twitter conditions (factor A).
$\bar{y}_{1 . .}=2.78 \quad \bar{y}_{2 . .}=2.32$
$\mathrm{SSA}=$ $\qquad$ $\mathrm{df}_{\mathrm{A}}=$ $\qquad$ MSA = $\qquad$
The sum of squares for sections (factor $B$ ) nested within Twitter conditions (factor $A$ ) is $\operatorname{SSB}(A)=5.00$, and the error sum of squares is $\operatorname{SSE}=130.5$.
p.11.b. Test $H_{0}: \alpha_{1}=\ldots=\alpha_{a}=0$

Test Statistic: $\qquad$ Rejection Region: $\qquad$ P-value < or > . 05
p.11.c. Test $H_{0}: \sigma_{b}^{2}=0 \quad H_{A}: \sigma_{b}^{2}>0$

Test Statistic: $\qquad$ Rejection

## Part F: Split-Plot Designs

QF.1. An experiment is conducted to compare five formulations of cookies and 4 cooking temperatures in an oven. Due to the nature of the experiment and time constraints, it was decided that on each of 7 days, there would be 4 cooking periods (one at each of the 4 temperatures), with each formulation being prepared in each cooking period. Give the Analysis of Variance table, including all sources of variation, degrees of freedom, and appropriate F-statistics. The response is a measure of cookie quality.

QF.2. A split-plot experiment is to be conducted to compare 4 grass varieties and 3 fertilizers on yield. Due to the nature of planting the grass varieties, they need to be planted on "large" strips of land, while fertilizer can be applied to "smaller" strips of land. Thus, a split-plot experiment will be conducted, with grass variety as the "whole plot" factor and fertilizer as the "subplot" factor. The experiment will be conducted on 5 strips (blocks) on a university's agricultural fields. The following model is to be fit (with grass and fertilizer as fixed factors, block as random):
$Y_{i j k}=\mu+\alpha_{i}+b_{j}+(a b)_{i j}+\gamma_{k}+(\alpha \gamma)_{i k}+\varepsilon_{i j k} \quad i=1, \ldots, 4 \quad j=1, \ldots, 5 \quad k=1,2,3$
$\sum_{i=1}^{4} \alpha_{i}=\sum_{k=1}^{3} \gamma_{k}=\sum_{i=1}^{4}(\alpha \gamma)_{i k}=\sum_{k=1}^{3}(\alpha \gamma)_{i k}=0 \quad b_{j} \sim N\left(0, \sigma_{b}^{2}\right) \quad(a b)_{i j} \sim N\left(0, \sigma_{a b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.2.a. You are given the following "schematic diagram" of the layout, as well as random numbers to be used for randomizing treatments to plots. Fill in which treatments would be assigned to each position, where A1, would represent Grass A/Fertilizer 1.

|  | Treatment |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Block1 |  |  |  |  |  |  |  |  |  |  |  |  |
| Block2 |  |  |  |  |  |  |  |  |  |  |  |  |
| Block3 |  |  |  |  |  |  |  |  |  |  |  |  |
| Block4 |  |  |  |  |  |  |  |  |  |  |  |  |
| Block5 |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | Grass |  |  |  |  |  |  |  |  |  |  |  |
| Block1 | 0.802 | 0.961 | 0.436 | 0.282 |  |  |  |  |  |  |  |  |
| Block2 | 0.042 | 0.038 | 0.205 | 0.803 |  |  |  |  |  |  |  |  |
| Block3 | 0.532 | 0.816 | 0.931 | 0.046 |  |  |  |  |  |  |  |  |
| Block4 | 0.721 | 0.919 | 0.312 | 0.715 |  |  |  |  |  |  |  |  |
| Block5 | 0.624 | 0.448 | 0.124 | 0.751 |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | Fertilizer |  |  |  |  |  |  |  |  |  |  |  |
| Block1 | 0.060 | 0.059 | 0.501 | 0.058 | 0.637 | 0.960 | 0.899 | 0.612 | 0.619 | 0.457 | 0.968 | 0.044 |
| Block2 | 0.431 | 0.594 | 0.141 | 0.881 | 0.546 | 0.298 | 0.077 | 0.307 | 0.408 | 0.761 | 0.157 | 0.830 |
| Block3 | 0.052 | 0.677 | 0.377 | 0.704 | 0.482 | 0.256 | 0.956 | 0.249 | 0.317 | 0.401 | 0.409 | 0.745 |
| Block4 | 0.743 | 0.972 | 0.385 | 0.040 | 0.309 | 0.535 | 0.051 | 0.543 | 0.585 | 0.507 | 0.287 | 0.738 |
| Block5 | 0.226 | 0.086 | 0.496 | 0.901 | 0.463 | 0.909 | 0.290 | 0.153 | 0.561 | 0.797 | 0.846 | 0.519 |

p.2.b. You are about to go into the field for a data collection mission. You decide to leave your advisor a set-up of the ANOVA in such a form that even he/she can't mess it up. Fill out following ANOVA table as simply as possible for him/her (assume he/she does know how to obtain the correct sums of squares):

| Source | df | SS | MS | F | F(.05) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Grass |  | SSG |  |  |  |
| Block |  | SSB |  |  | \#N/A |
| G*B |  | SSGB |  |  | \#N/A |
| Fert |  | SSF |  |  |  |
| G*F |  | SSGF |  |  |  |
| Error2 |  | SSE2 |  | \#N/A | \#N/A |
| Total |  | SSTot | \#N/A | \#N/A | \#N/A |

QF3. A split-plot experiment is to be conducted to compare 4 nitrogen sources and 3 time lengths of thatch accumulation on chlorophyll content $(Y)$ of grass. Due to the nature of planting the nitrogen sources, they need to be planted on "large" strips of grass, while time of thatch accumulation can be applied to "smaller" strips of grass. Thus, a split-plot experiment will be conducted, with nitrogen source as the "whole plot" factor and time of thatch accumulation as the "subplot" factor. The experiment will be conducted on 2 constructed putting greens (blocks) on a university's agricultural fields. The following model is to be fit (with nitrogen source and time of thatch accumulation as fixed factors, block as random). Also note, this is not repeated measures, as different sub-plots are observed at the 3 time lengths.
$Y_{i j k}=\mu+\alpha_{i}+b_{j}+(a b)_{i j}+\gamma_{k}+(\alpha \gamma)_{i k}+\varepsilon_{i j k} \quad i=1, \ldots, 4 \quad j=1,2 \quad k=1,2,3$
$\sum_{i=1}^{4} \alpha_{i}=\sum_{k=1}^{3} \gamma_{k}=\sum_{i=1}^{4}(\alpha \gamma)_{i k}=\sum_{k=1}^{3}(\alpha \gamma)_{i k}=0 \quad b_{j} \sim N\left(0, \sigma_{b}^{2}\right) \quad(a b)_{i j} \sim N\left(0, \sigma_{a b}^{2}\right) \quad \varepsilon_{i j k} \sim N\left(0, \sigma^{2}\right)$
p.3.a. You are given the following "schematic diagram" of the layout, as well as random numbers to be used for randomizing treatments to plots. Fill in which treatments would be assigned to each position, where A1, would represent Nitrogen A/Thatch 1.

| Treatmen | Cell 1 | Cell2 | Cell3 | Cell4 | Cell5 | Cell6 | Cell7 | Cell8 | Cell9 | Cell10 | Cell11 | Cell12 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Block1 |  |  |  |  |  |  |  |  |  |  |  |  |
| Block2 |  |  |  |  |  |  |  |  |  |  |  |  |


| Nitrogen | A | B | C | D |
| :--- | :---: | :---: | :---: | :---: |
| Block1 | 0.057 | 0.865 | 0.135 | 0.524 |
| Block2 | 0.340 | 0.514 | 0.198 | 0.807 |


| Thatch | 1 | 2 | 3 | 1 | 2 | 3 | 1 | 2 | 3 | 1 | 2 | 3 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Block1 | 0.004 | 0.631 | 0.759 | 0.939 | 0.175 | 0.216 | 0.070 | 0.959 | 0.025 | 0.259 | 0.831 | 0.457 |
| Block2 | 0.973 | 0.929 | 0.821 | 0.091 | 0.600 | 0.964 | 0.754 | 0.359 | 0.470 | 0.586 | 0.527 | 0.646 |

p.3.b. Complete the following ANOVA table:

| Source | df | SS | MS | F | F(.O5) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Nitrogen |  | 37.32 |  |  |  |
| Block |  | 0.51 |  | \#N/A | \#N/A |
| N*B |  | 1.26 |  |  | \#N/A |
| Thatch |  | 3.82 |  |  |  |
| N*T |  | 4.15 |  |  |  |
| Error2 |  | 48.78 | \#N/A | \#N/A | \#N/A |
| Total |  |  |  | \#N/A |  |

p.3.c. The p-values for testing for the various effects are:

Nitrogen: $>0.05$ or $<0.05$ Thatch: $>0.05$ or $<0.05$ Nit*Thatch: $>0.05$ or <0.05
p.3.d. Obtain Tukey's HSD for comparing all Nitrogen Sources, and for comparing all Thatch Time Lengths:
p.3.d.i. Nitrogen Sources:
p.3.d.ii. Thatch Time Lengths:

QF.4. A split-plot experiment is conducted to compare 4 cooking conditions (combinations of temperature/time) and 3 recipes for quality of taste of cupcakes. Because of the logistics of the experiment, each of the 4 cooking conditions can be conducted once per day (in random order). The recipes are randomly assigned to the slots in the oven (each recipe is observed once in each cooking condition). The experiment is conducted on 5 different days (blocks). Give the Analysis of Variance (sources and degrees of freedom and critical F-values), assuming no interaction between blocks and subplot units. The response is an average taste rating among a panel of judges.

| Source | Label | df | Error df | F(.05) |
| :--- | :--- | :---: | :---: | :---: |
| Whole Plot Factor |  |  |  |  |
| Blocks |  |  | \#N/A | \#N/A |
| Error1 |  |  | \#N/A | \#N/A |
| Sub Plot Factor |  |  |  |  |
| WP*SP Interaction |  |  |  |  |
| Error2 |  |  | \#N/A | \#N/A |
| Total |  |  | \#N/A | \#N/A |

QF.5. An ergonomic study was conducted as a Split-Plot design in Randomized Blocks. The response was lowest EMG signal for the Right Deltoid when a handwheel valve was being turned. The Whole-plot factor was the height of the handwheel with 4 levels (Overhead, Shoulder, Elbow, Knee). Blocks were 15 subjects who turned each handwheel at each angle. The Sub-Plot factor was angle with $\mathbf{3}$ levels $\left(90^{\circ}, 45^{\circ}, 0^{\circ}\right)$. Within each subject, the heights were assigned in random order, and each angle was measured in random order for that height. Complete the following ANOVA table.

| Source | df | SS | MS | F | F(.05) |
| :--- | ---: | ---: | :--- | :--- | :---: |
| Height(WP) |  | 11959 |  |  |  |
| Subj(BLK) |  | \#N/A | \#N/A | \#N/A | \#N/A |
| H $^{*}$ S(Err1) |  | 39651 |  | \#N/A | \#N/A |
| Angle(SP) |  | 10328 |  |  |  |
| H*A(WPxSP) | 117737 |  |  |  |  |
| Error2 | 112 | 91461 |  | \#N/A | \#N/A |
| Total | 179 | \#N/A | \#N/A | \#N/A | \#N/A |

The P-values for Interactions and Main effects are:

HeightxAngle: <0.05 or >0.05 Angle <0.05 or >0.05 Height <0.05 or >0.05

QF.6. An experiment was conducted as a Split Plot design in Randomized blocks. There were 3 recipes for cake mix, 4 cooking conditions (combinations of temperature/time), and the experiment was run on 5 days (blocks). The restriction on the randomization was that due to timing, only 4 cooking conditions could be conducted in a day. Thus, the 4 cooking conditions were randomly assigned to the order $1,2,3,4$ on a given day, and all recipes were cooked together. The recipes and cooking conditions are fixed factors, while day is considered random. The response measured was an overall quality rating from a panel of judges (average across judges).
$\qquad$ The Sub-Plot Factor? $\qquad$
p.6.b. Complete the following ANOVA Table.

| Source | SumSq | df | MS | F | F(.05) |
| :--- | ---: | :--- | :--- | :--- | :---: |
| WP | 679.1 |  |  |  |  |
| Block | 3291.3 |  |  | \#N/A | \#N/A |
| WP*Blk | 496.2 |  |  | \#N/A | \#N/A |
| SP | 1589.7 |  |  |  |  |
| WP*SP | 311.9 |  |  |  |  |
| Error | 731.9 |  |  | \#N/A | \#N/A |
| Total | 7100.1 | 59 | \#N/A | \#N/A | \#N/A |

p.6.c. Assuming the interaction is not significant, use Tukey's method to compare the 3 recipe means.

Recipe 1: $\bar{y}_{1}=56.19 \quad$ Recipe 2: $\bar{y}_{2}=54.06 \quad$ Recipe 3: $\bar{y}_{3}=44.36$
p.6.d. Assuming the interaction is not significant, use Tukey's method to compare the 4 cooking means.

Condition 1: $\bar{y}_{1}=56.65$ Condition 2: $\bar{y}_{2}=52.46$ Condition 3: $\bar{y}_{3}=48.72$ Condition 4: $\bar{y}_{4}=48.32$

QF.7. An grain study was conducted as a Split-Plot design in Randomized Blocks. The response was grain yield (kg/ha). The Whole-plot factor was the nitrogen fertilizing rate with $\mathbf{5}$ levels ( $0,45,90,135,180$ ). Blocks were $\mathbf{3}$ Years that the experiment was conducted in. The Sub-Plot factor was rice straw incorporation with $\mathbf{2}$ levels (absent, present).
p.7.a. Complete the following ANOVA table.

| ANOVA |  |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- | :--- |
| Source | df | SS | MS | F | F(.05) |
| WP (Nit) |  | 5693.63 |  |  |  |
| Block (Year) |  | 216.82 |  |  |  |
| WP*Block |  | 60.18 |  |  |  |
| SP (RiceStraw) |  | 110.82 |  |  |  |
| WP*SP |  | 0.96 |  |  |  |
| Error | 16.53 |  |  |  |  |
| Total |  | 6098.94 |  |  |  |

p.7.b. The P -values for Interactions and Main effects are:

HeightxAngle: $<0.05$ or $>0.05$ Angle $<0.05$ or $>0.05$ Height $<0.05$ or $>0.05$

| Nitrogen | Mean | Rice Straw | Mean |
| :---: | :---: | :---: | :---: |
| 0 | 48.65 | 0 | 73.02 |
| 45 | 75.19 | 1 | 76.87 |
| 90 | 79.07 |  |  |
| 135 | 85.85 |  |  |
| 180 | 85.96 |  |  |

p.7.c. Use Tukey's HSD to compare all pairs of Nitrogen levels on yield.
p.7.d. Obtain a 95\% CI for the difference in true means for Rice Straw Present - Absent

## Part G: Repeated Measures Designs

QG.1. A researcher is interested in comparing 4 diet plans. She selects 160 subjects and randomly assigns 40 subjects to each diet. She will measure their weight loss at 3 time points over the course of the year. Her analysis of variance will have the following sources of variation. Give her degrees of freedom for each source (actual numbers, not symbols)

| Source | Degrees of freedom |
| :--- | :--- |
| Diets |  |
| Subjects(Diet) --- Error1 |  |
| Time Points |  |
| Diets*Time |  |
| Time*Subjects(Diet) --- Error2 |  |
| Total |  |

QG.2. A repeated measures experiment was conducted to compare two treatments (zylkene and placebo) for cat anxiety. A total of 34 cats with anxiety were obtained, and randomized such that 17 received zylkene and 17 received placebo. Each cat was observed at 5 time points, and a global score of emotional state was observed (high scores are better). The following model is fit:

$$
y_{i j k}=\mu+\alpha_{i}+b_{j(i)}+\tau_{k}+(\alpha \tau)_{i k}+\varepsilon_{i j k} \quad i=1,2 j=1, \ldots, 17 k=1, \ldots, 5
$$

p.2.a. Complete the following ANOVA table.

| Source | df | SS | MS | F | F(.05) |
| :--- | ---: | ---: | :--- | :--- | :--- |
| Trt |  | 383 |  |  |  |
| Cat(Trt) |  | 2132 |  |  |  |
| Time |  | 324 |  |  |  |
| Time*Trt |  | 51 |  |  |  |
| Error2 |  |  |  |  |  |
| Total |  | 3278 |  |  |  |

p.2.a.i. The p-value for testing no time* treatment interaction is $<0.05$ or $>0.05$
p.2.a.ii. The $p$-value for testing no time main effect is
p.2.a.iii. The $p$-value for testing no treatment main effect is
$<0.05$ or $>0.05$
$<0.05$ or $>0.05$
p.2.b. Ignoring any potential interaction, obtain a 95\% Confidence Interval for the difference between Zyrtec and placebo effects (their means are 13.59 and 10.59, respectively).

QG.3. A repeated measures experiment was conducted to compare three battery recycling promotion strategies (humorous, factual, and control (no promotion)) for battery recycling ( $y=$ percent change from pre-ad recycling levels). A total of 21 stores were obtained, and randomized such that 7 received each strategy. Each store was observed at 8 time points, and a y was observed (negative scores imply lower levels than pre-treatment). The following model is fit:
$y_{i j k}=\mu+\alpha_{i}+b_{j(i)}+\tau_{k}+(\alpha \tau)_{i k}+\varepsilon_{i j k} \quad i=1,2,3 j=1, \ldots, 7 k=1, \ldots, 8 \quad$ with: $\quad \bar{y}_{1}=-7.4 \quad \bar{y}_{2}=35.8 \quad \bar{y}_{3}=-11.9, \quad \bar{y}=5.5$
p.3.a. Complete the following ANOVA table.

| Source | df | SS | MS | F | F(.05) |
| :--- | ---: | ---: | :--- | :--- | :--- |
| Strategy |  |  |  |  |  |
| Store(Strategy) |  | 10000 |  |  |  |
| Time |  | 10780 |  |  |  |
| Time*Strategy |  | 42000 |  |  |  |
| Error2 |  | 252000 |  |  |  |
| Total |  |  |  |  |  |

p.3.a.i. The p-value for testing no time*strategy interaction is $<0.05$ or $>0.05$
p.3.a.ii. The p-value for testing no time main effect is

$$
\begin{aligned}
& <0.05 \text { or }>0.05 \\
& <0.05 \text { or }>0.05
\end{aligned}
$$

p.3.a.iii. The $p$-value for testing no strategy main effect is
p.3.b. Ignoring any potential interaction, obtain Bonferroni's Minimum significant difference, and compare all pairs of strategies.

QG.4. A study compared doses of a drug on female rats' activity levels in a maze. A sample of 91 rats were selected, and randomized such that 21 rats received the Control Dose, 25 received Low Dose, 24 received Medium Dose, and 21 received High Dose. Each rat's activity levels were observed at 4 time points after dosing ( $15,30,45$, and 60 minutes). Hint: There are a total of $21+25+24+21=91$ rats in the study.
p.4.a. Complete the following ANOVA table.

| Source | df | SS | MS | F | F(.05) | Significant? |
| :--- | :--- | ---: | :--- | :--- | :--- | :---: |
| Dose |  | 16890 |  |  |  |  |
| Rat(Dose) |  | 176677 |  | \#N/A | \#N/A | \#N/A |
| Time |  | 176765 |  |  |  |  |
| Dose*Time |  | 2161 |  |  |  |  |
| Error2 |  | 92826 |  | \#N/A | \#N/A | \#N/A |
| Total | 363 | 465319 | \#N/A | \#N/A | \#N/A | \#N/A |

p.4.b. The means for each dose/time combination are given below. Use Bonferroni's method to compare all pairs of doses.

| Trt | n_dose | Time1 | Time2 | Time3 | Time4 | Mean |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Control | 21 | 165.9 | 136.3 | 121.4 | 103.4 | 131.8 |
| LowDose | 25 | 167.8 | 148.4 | 123.7 | 118.4 | 139.6 |
| MidDose | 24 | 168.3 | 140.3 | 117.2 | 109.8 | 133.9 |
| HighDose | 21 | 184.4 | 156.2 | 136.3 | 122.7 | 149.9 |
| Sum/Mean | 91 | 171.3 | 145.3 | 124.4 | 113.7 | 138.7 |

Note: When comparing doses $i$ and $j$, the standard error of the difference between dose means is:
$S E\left\{\bar{Y}_{i \bullet \bullet}-\bar{Y}_{j .}\right\}=\sqrt{M S_{\text {Sujjects }\left(T_{n s}\right)}\left(\frac{1}{n_{i} t}+\frac{1}{n_{j} t}\right)} \quad$ Values of $\left(\frac{1}{n_{i} t}+\frac{1}{n_{j} t}\right)$ are given below

| $\mathrm{i}, \mathrm{j}$ | 1,2 | 1,3 | 1,4 | 2,3 | 2,4 | 3,4 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $1 / \mathrm{tni}+1 / \mathrm{tnj}$ | 0.0219 | 0.0223 | 0.0238 | 0.0204 | 0.0219 | 0.0223 |

QG.5. A repeated measures design was used to compare the effects of Zylkene and Selgian Anipryl in dogs with anxiety disorders. There were 38 dogs, randomly assigned to the treatments ( 19 dogs per treatment). Each dog was measured at 5 time points, with a scale that has lower scores are better outcomes than higher scores).
p.5.a. Complete the following table.

| Source | df | SS | MS | F | F(.05) | Significant? |
| :--- | :--- | ---: | :--- | :--- | :---: | :---: |
| Treatment |  | 8.85 |  |  |  |  |
| Dog(Treatment) |  | 2020.42 |  | \#N/A | \#N/A | \#N/A |
| Time |  | 1573.45 |  |  |  |  |
| Time*Treatment |  | 20.44 |  |  |  |  |
| Error2 (Time*Dog(Trt)) |  | 1212.11 |  | \#N/A | \#N/A | \#N/A |
| Total |  | 4835.27 | \#N/A | \#N/A | \#N/A | \#N/A |

p.5.b. Assuming no significant interaction, obtain a $95 \%$ Confidence Interval for $\mu_{\mathrm{Z}}-\mu_{\mathrm{SA}}$. Note that the sample means are 19.99 and 19.56 for Zylkene and Selgian Anipryl, respectively.
p.5.c. Assuming no significant interaction, use Bonferroni's method to compare all pairs of Time Means. The sample means are: Time $_{1}=24.58 \quad$ Time $_{2}=21.21 \quad$ Time $_{3}=19.00 \quad$ Time $_{4}=17.63 \quad$ Time $_{5}=16.45$
$\begin{array}{lllll}\text { Time }_{5} & \text { Time }_{4} & \text { Time }_{3} & \text { Time }_{2} & \text { Time }_{1}\end{array}$
QG.6. A repeated measures design was conducted to compare 3 treatments for dry skin (Placebo, WPLC-O, WPLC-P). The study had a total of 60 subjects who were randomized so that 20 subjects received Placebo, 20 received WPLC-O, and 20 received WPLC-P. Each subject was measured on 3 days (Days 15, 30, 60) for the respons skin hydration.
p.6.a. Complete the following ANOVA table.

| Source | df | SS | MS | $F$ | F(.05) |
| :--- | ---: | ---: | :--- | :--- | :--- |
| Trts |  | 3475 |  |  |  |
| Subj(Trt) |  | 3245 |  | \#N/A | \#N/A |
| Time |  | 575 |  |  |  |
| TrtxTime |  | 128 |  |  |  |
| Error2 | 114 | 4868 |  | \#N/A | \#N/A |
| Total | 179 | 12291 | \#N/A | \#N/A | \#N/A |

p.6.b Are the following effects significant? TrtxTime Interaction $\qquad$ Trts $\qquad$ Time $\qquad$
p.6.c. The means for the Treatments are: Placebo: 47.4 WPLC-O: 57.3 WPLC-P: 55.9. Use Bonferroni's method to compare all pairs of treatments (assuming no interaction).
$\qquad$

QG.7. A study was conducted among obese Thai subjects on the effect of drinking green tea on weight. There were a total of 60 subjects, and were randomized so that 30 drank green tea and 30 received a placebo. Each subject's weight was measured in kilograms at 4,8 , and 12 weeks after intervention. The model was conducted as a Repeated Measures design with 2 treatments, 30 subjects per treatment, and 3 time points.
p.7.a. Complete the following ANOVA table.

| Source | df | SS | MS | F | F(.05) |
| :--- | :--- | :--- | :--- | :--- | :---: |
| Treatment |  | 617.1605 |  |  |  |
| Subject(Trt) |  | 11388.05 |  | \#N/A | \#N/A |
| Time |  | 67.357 |  |  |  |
| Trt*Time |  | 47.221 |  |  |  |
| Error2 |  | 3796.017 |  | \#N/A | \#N/A |
| Total |  | 15915.81 | \#N/A | \#N/A | \#N/A |

p.7.b. Is there a significant treatment by time interaction? Yes / No
p.7.c. Is there a significant treatment main effect? Yes / No
p.7.d. Is there a significant time main effect? Yes / No
p.7.e. The mean weights for the green tea and placebo groups across time points are $\bar{y}_{g}=66.34 \quad \bar{y}_{p}=70.04$

Compute a $95 \%$ Confidence Interval for the difference in their effects on weight.

## Part H: General Design Identification Problems

QH.1. For the following problems, identify the factor(s), state whether they are fixed or random, give the analysis of variance table, with sources of variation and degrees of freedom, symbolic F-ratios and critical values (just give degrees of freedom) for relevant significance tests for treatment factors.
p.1.a. An ergonomic study was conducted to compare 6 car seat designs in terms of an overall comfort index (Y). A sample of 12 subjects was selected, and each subject rated each car seat one time.
p.1.b. A food preference study was interested in the main effects and interactions among two factors on subjects' ratings of attractiveness of a plate of food. The factors under study were plate color (monochrome, color) and balance of food placement on the plate (symmetric (balanced), asymmetric (unbalanced)). A sample of 68 subjects were selected, and randomized such that 17 received each combination of color and balance. Each person only rated one plate.
p.1.c. A study was conducted to compare 6 models of bread machines on quality of baked bread. There were 6 varieties of bread, and 6 chefs, and each variety was made by each machine once, and each chef used each machine once. The response was an overall quality rating based by a panel of judges (which was combined to a single rating).
p.1.d. A study was conducted to measure the reliability of collegiate gymnastics judges, and variation in gymnast skills. A sample of 8 judges was selected, and a sample of 4 gymnasts was selected. Each gymnast was filmed on 3 occasions, and each judge rated the 3 videos.

QH.2. For the following problems, identify the factor(s), state whether they are fixed or random, give the analysis of variance table, with sources of variation and degrees of freedom, symbolic F-ratios and critical values for $\alpha=0.05$ significance level tests.
p.2.a. A researcher is interested in studying the variation in laboratories measuring the levels of nutrient in batches of raw materials. Her department has contracts with 6 laboratories, and she obtains a random sample of 5 batches of the raw material, dividing each batch into 24 sub-batches. She sends each laboratory 4 randomly chosen sub-batches from each of the 5 batches and has each laboratory measure the nutrient levels in each of their 20 assigned sub-batches.
p.2.b. A study measured emulsion properties when different types of plant oils (soybean, hazelnut, canola, sunflower, corn, cotton, and olive) are applied to different types of meat (chicken, beef, and turkey). Each type of oil was applied to each type of meat, and there were 2 replicates per treatment (combination of plant oil and meat type).
p.2.c. An experiment is conducted in a field to measure the effects among 5 seeding rates in an experimental field that is set up on plots set in a $5 \times 5$ array with 5 rows and 5 columns. The rates are applied to the field such that each rate is applied once on each row and once on each column.
p.2.d. A study compares 4 popular diets on weight loss. A sample of 160 overweight subjects are obtained and assigned at random, such that each diet has 40 subjects. Weight loss over 30 days is measured.

QH.3. For the following problems, identify the factor(s), state whether they are fixed or random, give the analysis of variance table, with sources of variation and degrees of freedom, symbolic F-ratios and critical values for $\alpha=0.05$ significance level tests.
p.3.a. An experiment was conducted to determine the effects of 2 factors on growth of cucumbers in a greenhouse. Factor A was irrigation method (Furrow and Subsurface Drip) and Factor B was post-irrigation aereation level $(0,0.50,0.75,1)$. Due to the nature of the experimental set-up, the experiment was conducted as a Split-Plot Design with irrigation method as the Whole-plot factor, and aereation level as the Sub-plot factor. The experiment was conducted in 5 blocks. The response was cucumber length.
p.3.b. A steel experiment was conducted to compare 3 levels of titanium carbon content. Two blocking factors were included: sliding velocity and applied load (each with 3 levels). Each titanium carbon content was applied to each sliding velocity and to each applied load once. The response was wear rate.
p.3.c. A study compared the swimming speeds of male and female zebra-fish at 4 rearing temperatures. The experimenters raised 48 males and 48 females, 12 of each gender at each of the 4 temperatures ( $22 \mathrm{C}, 25,28$, and 31 ). The response is relative critical swimming speed.

QH.4. For the following scenarios, give the sources of variation, their degrees of freedom, appropriate F-ratios, and critical values for the F-tests.
p.4.a. A mock jury experiment was conducted among college students to study the effects of 2 factors on judgments of witness effectiveness (Y). The factors were: Defendant's group identity (Factor A: In-group (similar background to students), Neutral (no information given), Out-Group (member of a radical political group)) and Eyewitness Testimony (Factor B: Consistent during cross-examiniation, Inconsistent). There were a total on $\mathrm{N}=180$ students in the experiment, with them randomized so that 30 were assigned to each of 6 combinations of levels of Factors A and B in a Completely Randomized Design (each subject was in exactly one condition).
Source $\quad$ DF F-ratio $\quad F(.05)$
p.4.b. An engineer is interested in variation in products and operator measurements in a factory. She samples 5 parts and 4 employees within her factory and has each engineer measure each part 3 times, in random order. The response is the measurement of the part.

Source
DF F-ratio
$F(.05)$
p.4.c. A large juice producer makes 3 varieties of juice (Orange, Lemon, and Apple). For each variety, the company has many farms that provide fruit of that specific variety (no farm produces fruits of more than one variety). They are interested in measuring sugar concentration in fruit (Y). Random samples of 4 farms are selected from each variety, with 20 fruits being sampled from each farm, and sugar concentration is measured on each fruit.

Source
DF
F-ratio
$F(.05)$

